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# Demand and supply of differentiated products

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#### • Reviews:

- Gandhi and Nevo (2021)
- Berry and Haile (2021)
- Aguirregabiria (2021) chapter 2
- Hortaçsu and Joo (2023) 2.1-2.2 and chapter 3
- Ackerberg et al. (2007) section 1 (these slides use their notation)
- Reiss and Wolak (2007) sections 1-7, especially 7
- Classic papers:
  - Berry (1994)
  - Berry, Levinsohn, and Pakes (1995)

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Section 1

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#### Introduction

- Typical market for consumer goods has many differentiated, but similar products, e.g.
  - Cars
  - Cereal
- Differentiated products are a source of market power
- Having many products can result in many parameters creating estimation difficulties and requiring departures from textbook demand and supply models

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#### Motivation

- Counterfactuals that do not change production technology
  - Mergers
  - Tax changes
- Effects of new goods
- Cost-of-living indices
- Product differentiation and market power
  - Cross-price elasticities

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#### Section 2

# Demand in product space

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# Demand in product space 1

- J products, each treated as separate good
- Classical demand,

$$q_1 = D_1(p_1, ..., p_J, z_1, \eta_1; \beta_1)$$
  
 $\vdots = \vdots$   
 $q_J = D_J(p_1, ..., p_J, z_J, \eta_J; \beta_J),$ 

and supply (firms' first-order conditions for prices):

$$p_1 = g_1(q_1, ..., q_J, w_1, v_1; \theta_1)$$

$$\vdots = \vdots$$

$$p_J^d = g_J(q_1, ..., q_J, w_J, v_J; \theta_J),$$

where

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### Demand in product space 2

- $p_i = price$
- $q_j = quantity$
- $z_j$  = observed demand shifter
- $\eta_i$  = unobserved demand shock
- $\beta_i$  = demand parameters
- $w_i$  = observed supply shifter
- $v_i$  = unobserved supply shock
- $\theta_i$  = supply parameters
- $D_j$  typically parametrically specified, e.g.

$$\ln q_j = \beta_{j0} + \beta_{j1}p_1 + \dots + \beta_{jJ}p_J + \beta_{jy}\ln y + Z_1\gamma + \nu_j$$

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# Demand in product space

Use reduced form to find instruments

$$q_{1} = \prod_{1}^{q}(Z, W, v, \eta; \beta, \theta)$$

$$\vdots = \vdots$$

$$q_{J} = \prod_{J}^{q}(Z, W, v, \eta; \beta, \theta)$$

$$p_{1} = \prod_{1}^{p}(Z, W, v, \eta; \beta, \theta)$$

$$\vdots = \vdots$$

$$p_{J} = \prod_{J}^{p}(Z, W, v, \eta; \beta, \theta)$$

- Cost shifters of product *j* excluded from demand and supply of product *k*, but in reduced form
  - Cost data often not available
  - If available, unlikely to be product specific
- Attributes of other products
  - Hausman (1996) uses prices of other products
  - Hard to justify, especially with prices

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### Demand in product space 1

- Advantages of product space:
  - Flexible substitution patterns
  - Does not require detailed product attribute data
- Problems with product space:
  - Representative agent and aggregation issues
    - With heterogeneous preferences, aggregate market demand need not meet restrictions on individual demand derived from economic theory
    - Cannot use restrictions easily to improve estimates
    - Can use simulation to aggregate (Pakes, 1986)
  - 2 Too many parameters,  $O(J^2)$ 
    - Can limit by restricting cross-price elasticities, e.g. Pinkse, Slade, and Brett (2002)
  - 3 Too many instruments needed, J
  - Cannot analyze new goods

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#### Section 3

# Demand in characteristic space

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# Demand in characteristic space

Why do firms differentiate products?

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# Demand in characteristic space

- Why do firms differentiate products?
- Because consumers have heterogeneous tastes for product characteristics
  - E.g. cars: tastes for size, safety, fuel efficiency, etc

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### Demand in characteristic space

 Model consumer preferences for characteristics and treat products as bundles of characteristics

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### Demand in characteristic space

- Model consumer preferences for characteristics and treat products as bundles of characteristics
- Reduces number of parameters
- · Predict demand for new goods
- Demand system consistent with utility maximization

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#### Demand in characteristic space

- Model consumer preferences for characteristics and treat products as bundles of characteristics
- Reduces number of parameters
- Predict demand for new goods
- Demand system consistent with utility maximization
- Early work: Lancaster (1971), McFadden (1973)
- Key extension to early work: Berry, Levinsohn, and Pakes (1995)

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Early work

# Early work in characteristic space

- Consumer chooses one or none of J products
- Utility of consumer i from product i

$$u_{ij} = x_j \beta + \epsilon_{ij}$$

with  $\epsilon_{ii}$  iid across i and j (usually Type I extreme value)

Implies aggregate demand (for Type I extreme value)

$$q_j = \frac{\exp(x_j \beta)}{1 + \sum_{k=1}^{J} \exp(x_k \beta)}$$

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### Downside of Logit

- Problem: restrictive substitution "independence of irrelevant alternatives"
  - Two goods with the same shares have the same cross price elasticities with any third good (think about a luxury and bargain good with equal shares)
  - Goods with same shares should have same markups
- Solution: add heterogeneity in  $\beta$  and/or allow correlation across j in  $\epsilon_{ij}$

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#### Model 1

- Consumers i, goods j, markets t
- Utility: (include good 0 = buy nothing)

$$u_{ijt} = U \begin{pmatrix} p_{roduct \ characteristics} \\ p_{jt}, \tilde{x}_{jt}, & \xi_{jt} \\ observed \ unobserved \end{pmatrix} \xrightarrow{consumer \ characteristics} z_{it}, v_{it} \\ observed \ unobserved \end{pmatrix}$$

- $x_{it} = (\tilde{x}_{it}, p_{it}) \in \mathbb{R}^K$ ,  $z_{it} \in \mathbb{R}^R$ ,  $v_{it} \in \mathbb{R}^L$
- Choose *j* if  $u_{ijt} > u_{ikt} \ \forall k \neq j$

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#### Model 1

• Usually *U*(⋅) linear:

$$u_{ijt} = \underbrace{x_{jt}}_{1 \times K} \underbrace{\theta_{it}}_{\theta_{it}} + \underbrace{\xi_{jt}}_{1 \times 1} + \epsilon_{ijt}$$

$$= \theta + \theta^{o} z_{it} + \theta^{u} v_{it}$$

for j = 1...J and normalize  $u_{i0t} = 0$ 

- Assume  $\epsilon_{iit}$  i.i.d. double exponential
- Assume  $v_{it} \sim f_{\nu}(\cdot; \theta)$ , e.g. independent normal
- Write as product specific + observed interactions + unobserved interactions

$$u_{ijt} = \underbrace{\delta_{j}}_{=x_{it}\bar{\theta} + \xi_{jt}} + x_{jt} \underbrace{\theta^{o}}_{K \times R} z_{it} + x_{jt} \underbrace{\theta^{u}}_{K \times L} v_{it} + \epsilon_{ijt}$$

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#### Endogeneity

- Usually assume  $E[v_{it}|x_{jt},z_{it}]=0$  and  $E[\epsilon_{ijt}|x_{jt},z_{it}]=0$ 
  - Not interested in counterfactuals with respect to changes in  $z_{it}$ , so can treat as residual, i.e.

$$v_{it} = \theta_{it} - \mathsf{E}[\theta_{it}|z_{it}]$$

• Market average  $v_{it}$  or  $\epsilon_{ijt}$  plausibly correlated with  $p_{jt}$  or other product characteristics, but this correlation absorbed into  $\xi_{jt}$  and/or market fixed effects

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#### Endogeneity

- Problem is  $\xi_{jt}$ 
  - Prices and other flexible product characteristics will be correlated with ξ<sub>jt</sub>
  - If  $\xi_{jt}$  serially correlated, then possibly also correlated with inflexible product characteristics
  - Need instrument(s),  $w_{it}$  such that  $E[\xi_{it}|w_{it}] = 0$ 
    - Cost shifters
    - Characteristics of other products

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# Aggregate data 1

- Often only have data on product characteristics and market shares
- Maybe also distribution of some individual characteristics for each market (e.g. income and education from CPS or census)
- Instrument w such that  $E[\xi_j|w] = 0$
- Distribution of  $v \sim f_v(\cdot; \theta_v)$ 
  - Combination of estimated market level distribution of observed individual characteristics and parametric distributions of unobserved individual characteristics
  - e.g.  $v_{it} = (educ_{it}, income_{it}, e_{it})$

$$F_{v,t}(s, y, e; \theta_v) = \underbrace{\hat{F}_t(s, y)}_{\text{empirical distribution}} \Phi\left(\frac{e - \theta_v^{\mu}}{\theta_v^{\sigma}}\right)$$

 $\hat{F}_t(s, y)$  estimated from CPS or other similar data set

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### Aggregate data 2

- Assume  $\epsilon_{ijt} \sim$  double exponential (aka Gumbel or type I extreme value) as in logit
  - Computationally convenient, but other distributions feasible too

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#### Estimation outline

• Estimate  $\theta$  from moment condition

$$\mathsf{E}[\xi(\cdot;\theta)|w]=0$$

- Where ξ(·; θ) is such that model predicted market shares = observed market shares<sup>1</sup>
  - **1** Compute shares given  $\theta$ ,  $\sigma(\cdot; \theta, \delta)$
  - 2 Find  $\delta(\cdot; \theta) = x_{jt}\overline{\theta} + \xi(\cdot; \theta)$  such that observed shares,  $s_{jt}$  = model shares,  $\sigma(\cdot; \theta, \delta)$ , then

$$\xi(\cdot;\theta) = \delta(\cdot;\theta) - x_{jt}\bar{\theta}$$

 $<sup>^{1}</sup>$ In this slide  $\cdot$  means the data. I will leave the  $\cdot$  out of the notation in subsequent slides. I will also leave out t subscripts.

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## Computing model shares

• Integrate over ν

$$\sigma_{j}(\theta, \delta) = \int \frac{\exp(\delta_{j} + x_{j}\theta^{u}v)}{1 + \sum_{k=1}^{j} \exp(\delta_{k} + x_{k}\theta^{u}v)} dF_{v}(v)$$

 Integral typically has no closed form, so compute numerically, usually by Monte Carlo integration

$$\sigma_{j}(\theta, \delta) = \sum_{r=1}^{N_{s}} \frac{\exp(\delta_{j} + x_{j}\theta^{u}v_{r})}{1 + \sum_{k=1}^{j} \exp(\delta_{k} + x_{k}\theta^{u}v_{r})}$$

where  $v_r$  are  $N_s$  random draws from  $f_v$ 

- Issues about how best to compute integral simulation vs quadrature, type of simulation (Skrainka and Judd, 2011)
- Simulation (more generally approximation) of integral affects distribution of estimator

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# Solving for $\delta$ and $\xi$

- Want  $\delta$  s.t.  $\sigma_j(\theta, \delta) = \hat{s}_j$
- Berry, Levinsohn, and Pakes (1995) show

$$T(\delta) = \delta + \log(\hat{s}_j) - \log(\sigma_j(\theta, \delta))$$

#### is a contraction

- Unique fixed point  $\delta$  such that  $\delta = \delta + \log(\hat{s}_j) \log(\sigma_j(\theta, \delta))$ , i.e.  $\hat{s}_j = \sigma_j(\theta, \delta)$
- Can compute  $\delta(\theta)$  by repeatedly applying contraction (in theory and practice often faster to use other method)
- $\xi_j(\theta) = \delta_j(\theta) x_j \bar{\theta}$
- Important identifying assumption: only  $\theta$  s.t.  $\xi_j(\theta) = \xi_j^0$  is true  $\theta_0$

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# Estimating $\theta$

- Conditional moment restriction  $E[\xi_j(\theta)|w] = 0$
- Empirical unconditional moments:

$$G_{J,T,N,N_s} = \frac{1}{JT} \sum_{j=1}^{J} \sum_{t=1}^{T} \xi_{jt}(\theta) f(w_t)$$

#### where

- f(w) = vector of function of w
- *J* = number of products
- *T* = number of markets
- N = number of observations in each market underlying  $\hat{s}_i$
- $N_s$  = number of simulations
- Asymptotic properties (consistency, distribution), depend on which of J, T, N, and  $N_s$  are  $\to \infty$ , see Berry, Linton, and Pakes (2004)
- Reynaert and Verboven (2014): using optimal instruments greatly improves efficiency and stability

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## Pricing equation 1

- More moments give more precise estimates
- Assumption about form of equilibrium allows use of firm first order condition (pricing equation) as additional moment
- Nash equilibrium in prices
- · Log linear marginal cost

$$\log mc_j = r_j \theta^k + \omega_j$$

- r<sub>j</sub> = observed product characteristics, input prices, maybe quantity, etc
- $\omega_j$  = unobserved productivity, possibly endogenous
- Firm f producing set of product  $\mathcal{J}_f$ ,

$$\max_{p_j:j\in\mathcal{J}_f}\sum_{j\in\mathcal{J}_f}\left(p_j-C_j(\cdot)\right)Ms_j(\cdot,p)$$

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## Pricing equation 2

First order condition:

$$\sigma_j(\cdot) + \sum_{l \in \mathcal{J}_f} (p_l - mc_l) \frac{\partial \sigma_l(\cdot)}{\partial p_j} = 0$$

Collect as

$$s + (p - mc)\Delta = 0$$

· Rearrange and use log linear marginal cost

$$\log(p - \Delta^{-1}\sigma) - r\theta^{c} = \omega(\theta)$$

- Conditional moment restriction  $E[\omega(\theta)|w] = 0$
- Add empirical moments to G,  $\frac{1}{JT} \sum_{jt} \omega_{jt}(\theta) f(w_t)$

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#### Micro data

- Berry, Levinsohn, and Pakes (2004)
- Data on individual choices and characteristics

$$u_{ijt} = \underbrace{\delta_{j}}_{=x_{jt}\bar{\theta} + \xi_{jt}} + x_{jt} \underbrace{\theta^{o}}_{K \times R} z_{it} + x_{jt} \underbrace{\theta^{u}}_{K \times L} v_{it} + \epsilon_{ijt}$$

- Random coefficients discrete choice model, so can identify and estimate  $\delta$ ,  $\theta^o$ , and  $\theta^u$  without assumptions about  $\xi$  and x
  - Ichimura and Thompson (1998) give conditions for nonparametric identification of random coefficients binary choice models
  - · Estimate by MLE or (usually) GMM
- Still need  $\bar{\theta}$  for price elasticities, etc

$$\delta_j = x_{jt}\bar{\theta} + \xi_{jt}$$

- Use IV
- Use IV with a pricing equation

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