

Dynamic Oligopoly: Additional Issues

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References

- **Reviews:**
 - Aguirregabiria and Nevo (2010)
 - Aguirregabiria (2017) chapters
 - Akerberg, Caves, and Frazer (2015) section 3
 - Aguirregabiria and Mira (2010)

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Section 1

Introduction

Introduction

- There are not many applied papers that estimate dynamic games (less true now than 4 years ago)
- Reasons:
 - ① Estimating dynamic games is computationally intensive
 - ② Assumption that the only unobserved heterogeneity are i.i.d shocks is not plausible
 - Why bother estimating a complicated model if the results are not credible?
 - Should add some permanent and/or autocorrelated unobserved heterogeneity
- Today we will look at recent research addressing these two issues

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Section 2

Computation

- Estimation involves maximizing some objective function subject to equilibrium conditions
- Estimation methods:
 - Maximum likelihood

$$\max_{\theta \in \Theta, \mathbf{P} \in [0,1]^N} \sum_{m=1}^M \sum_{t=1}^{T_m} \sum_{i=1}^N \log \Lambda(a_{imt} | v_i^{\mathbf{P}}(\cdot, x_{mt}; \theta))$$

$$\text{s.t. } \mathbf{P} = \Lambda(v^{\mathbf{P}}(\theta))$$

- 2-step estimators: estimate $\hat{P}(a|x)$ from observed actions and then

$$\max_{\theta \in \Theta} \sum_{m=1}^M \sum_{t=1}^{T_m} \sum_{i=1}^N \log \Lambda(a_{imt} | v_i^{\hat{P}}(\cdot, x_{mt}; \theta))$$

Computation 2

- Nested pseudo likelihood (NPL) (Aguirregabiria and Mira, 2007): after 2-step estimator update

$\hat{\mathbf{P}}^{(k)} = \Lambda(v^{\hat{\mathbf{P}}^{(k-1)}}(\hat{\theta}^{(k-1)}))$, re-maximize pseudo likelihood to get $\hat{\theta}^{(k)}$ and repeat

- Computation time: 2-step $<$ NPL \leq MLE
- Possible reductions in computation
 - Improve calculation of $\Lambda(v^{\mathbf{P}}(\theta))$ (main problem is $v^{\mathbf{P}}$)
 - Improve maximization
 - Better maximization algorithm (MPEC Su and Judd (2012)); same issues with starting values and local optima as in BLP models
 - Bayesian method (MCMC) instead of maximization, Imai, Jain, and Ching (2009) and Gallant, Hong, and Khwaja (2012)

Improving calculation of value function 1

- For finite state space can compute v^P as

$$V^P(\theta) = (I - \delta \mathbf{M}_c)^{-1} \mathbf{M}_c [\pi(\theta) + g(\mathbf{p}, \theta)]$$

- Inverting matrix takes $O(S^3)$ operations where S is size of state space
 - Matrix inversion becomes prohibitively slow for surprisingly moderate S
 - On my desktop (AMD-FX8150 cpu) $S = 1000$ takes 1.15 seconds, $S = 2000$ takes 11.1 seconds, $S = 3000$ takes 38.6 seconds, $S = 4000$ takes 91.2 seconds
 - Faster hardware can cut these times by a constant factor, but still face cubic growth
 - Using GPU instead of CPU for matrix inversion can be much faster

Improving calculation of value function 2

- On my desktop the GPU (NVidia GeForce GTX 560) takes about a hundredth as long to invert large matrices as the CPU
- Scientific computing using GPUs is a new and active field
- Programming for GPUs can be difficult
- Fast GPUs are not part of most servers
- Inverting sparse matrices can take much less than $O(S^3)$ operations
 - $I - \delta \mathbf{M}_c$ is often sparse
 - Exact complexity of inversion depends on number of non-zero entries and their locations (sparsity pattern)
- Some papers iterate value equation instead of explicitly inverting

$$V^p(\theta) = \mathbf{M}_c [\pi(\theta) + g(\mathbf{p}, \theta) + \delta V^p(\theta)]$$

- Simulation often used

Improving calculation of value function 3

- Still solving same equation, if solving accurately has to take $O(S^3)$
 - If iterating is faster must be either (i) implicitly exploiting sparsity or (ii) solving inaccurately
 - Estimation can proceed with approximate solutions that only become exact at estimated θ , e.g. [Kasahara and Shimotsu \(2011\)](#)
- **M** depends on **P**, so for 2-step methods only need to compute inverse once
 - State space can be very large even for models that appear simple
 - E.g. entry/exit game with N firms whose identities matter $S = 2^N |X|$

Reducing the size of the state space 1

- Economically motivated restrictions can reduce the size of the state space

R1 assume homogenous players and symmetric equilibrium

- E.g. entry game, assume:
 - 1 Only number of competitors and not their identities affects profits
 - 2 Firms have the same profit function
 - 3 Symmetric equilibrium

then state space size is $2(N + 1)|X|$

R2 Inclusive values (Nevo and Rossi, 2008)

- Inclusive value = in discrete choice model the expected utility of a consumer from facing several options before observing the shocks (McFadden et al., 1978)

$$E \left[\max_j u_j + \epsilon_j \right]$$

Reducing the size of the state space 2

- Adjusted inclusive value \approx inclusive value minus firm's marginal costs; denote by i_f
- With appropriate assumptions, profits can be written as function of adjusted inclusive values,

$$\pi_f(\text{all state variables}) = \pi_f(i_f, i_{-f})$$

- Assume strategies only depend on adjusted inclusive values,

$$P(i_{f,t+1}, i_{-f,t+1} | \text{state}_t) = P(i_{f,t+1}, i_{-f,t+1} | i_{f,t}, i_{-f,t}, \sum \text{invest}_{f,t})$$

possible justifications:

- Strong assumptions about investment process
- Limited information of firms
- Bounded rationality: firms have as hard a time computing strategies as we do

Reducing the size of the state space 3

- Then value function only depends on inclusive values
- R3 Oblivious equilibrium (Weintraub, Benkard, and Van Roy (2008), Weintraub, Benkard, and Van Roy (2010), Farias, Saure, and Weintraub (2012))
- Oblivious equilibrium: firms make decisions conditional only on their own state variables and long-run industry average state
 - In Markov equilibrium, firms make decisions based on all state variables
 - Weintraub, Benkard, and Van Roy (2008) show that oblivious equilibrium approximates Markov perfect equilibrium as number of firms increases
 - Krusell and Smith (1998) use similar idea in dynamic macro model

- Value function:

$$\begin{aligned}V^P(\theta) &= \mathbf{M}_c [\pi(\theta) + g(\mathbf{p}, \theta) + \delta V^P(\theta)] \\ &= \mathbf{M}_c [\pi(\theta) + g(\mathbf{p}, \theta) + \delta \mathbf{M}_c [\pi(\theta) + g(\mathbf{p}, \theta) + \delta V^P(\theta)]] \\ &= \left(\sum_{t=0}^T \mathbf{M}_c^{t+1} \delta^t \right) [\pi(\theta) + g(\mathbf{p}, \theta)] + \mathbf{M}_c^{T+2} \delta^{T+1} V^P(\theta)\end{aligned}$$

- **Arcidiacono and Miller (2011)**: if $\mathbf{M}_c^{T+2} \delta^{T+1} V^P(\theta)$ is identical across actions then it will drop out of $v^p(a, x) - v^p(a', x)$, avoiding inversion
- Examples:
 - Renewal action: bus engine replacement
 - Terminal choice

Section 3

Unobserved heterogeneity

Unobserved heterogeneity 1

- Only unobservables in basic dynamic model are i.i.d. shocks
- More plausible to allow richer unobserved heterogeneity i.e. unobserved state variables
- Panel data can identify fairly rich unobserved heterogeneity
 - E.g. in linear models:
 - Random effects
 - Fixed effects
 - (Dynamic factors)
 - Fixed effects with additional autocorrelation (dynamic panel models as in [Blundell and Bond \(2000\)](#))
 - Variants of methods for linear models can be applied to dynamic games, but not straightforward because of
 - Nonlinearity – requires different identification arguments; complicates fixed effects estimation

Unobserved heterogeneity 2

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- Computation – introducing unobserved state variables makes computing the model more complex
- Types of unobserved heterogeneity
 - Permanent firm or market unobserved heterogeneity
 - Similar to random or fixed effects
 - Aguirregabiria and Mira (2007), Collard-Wexler (2013), Aguirregabiria and Ho (2012)
 - Unobserved states that follow a controlled Markov process
 - Identification of transition probabilities: Kasahara and Shimotsu (2009), Hu and Shum (2012), Allman, Matias, and Rhodes (2009), Hu and Shum (2013)
 - Arcidiacono and Miller (2011), Kasahara and Shimotsu (2011)

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- Here we go over approach of Aguirregabiria and Mira (2007) as described in Aguirregabiria and Nevo (2010)
- Market specific random effect in profits

$$\pi_{imt} = \Pi(a_{mt}, x_{imt}, \theta) + \theta_i(a) + \sigma_i \xi_m + \epsilon_{imt}$$

- ϵ_{imt} i.i.d.
- ξ_m unobserved, discrete with finite support, known mean and variance (absorbed by θ_i and σ_i), known support, $\{\xi^\ell\}_{\ell=1}^L$, pmf λ
- θ_i and σ_i varying with i requires large T (or large M and same firms across markets)
- Conditional choice probabilities different for each ℓ , denote by \mathbf{P}_ℓ

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- Equilibrium for each ℓ :

$$\mathbf{P}^\ell = \Lambda(\mathbf{v}^{\mathbf{P}^\ell}(\theta), \ell)$$

- Pseudo-likelihood integrates over distribution of ξ_m

$$\begin{aligned} \max_{\theta, \mathbf{P}^\ell, \lambda} & \sum_{m=1}^M \sum_{t=1}^T \sum_{i=1}^N \log \left(\sum_{\ell=1}^L \lambda_{\ell|x} \Lambda \left(\mathbf{a}_{imt} | \mathbf{v}_i^{\mathbf{P}^\ell}(\cdot, \mathbf{x}_{mt}; \theta, \ell) \right) \right) \\ \text{s.t. } & \mathbf{P}^\ell = \Lambda(\mathbf{v}^{\mathbf{P}^\ell}(\theta), \ell) \end{aligned}$$

where

$$\lambda_{\ell|x} = P(\xi_m = \xi^\ell | \mathbf{x}_{m1})$$

- Initial conditions problem

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- ξ_m will be correlated with initial values of endogenous state variables (markets with high ξ_m will start with a large number of firms)
- One solution: assume stationary, find stationary distribution of $x|\xi$,

$$P(x_t|\xi_m = \xi^\ell) = \sum_{x_{t-1}} P(x_{t-1}|\xi_m = \xi^\ell)P(x_t|x_{t-1}, \xi_m = \xi^\ell)$$

Bayes' rule

$$\lambda_{\ell|x} = \frac{\lambda_\ell P(x|\xi_m = \xi^\ell)}{\sum_{j=1}^L \lambda_j P(x|\xi_m = \xi^j)}$$

- To apply 2-step estimators need to first consistently estimate transition probabilities conditional on unobserved ξ_m and $P^\ell(\cdot|x_{mt})$

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- Can use [Kasahara and Shimotsu \(2009\)](#), but [Aguirregabiria and Nevo \(2010\)](#) say estimation is difficult
- Identification argument is constructive, based on singular value decomposition, can mimic for estimation, e.g. [Hu, Shum, and Tan \(2010\)](#)
- [Levine, Hunter, and Chauveau \(2011\)](#)
- NPL can be used but must iterate to convergence unless started from consistent $\hat{P}^\ell(\cdot|x_{mt})$
 - Start with arbitrary $P^\ell(\cdot|x_{mt})$
 - Maximize pseudo likelihood to get $\hat{\theta}$, distribution of ξ_m
 - Update $P^\ell(\cdot|x_{mt})$
 - Repeat until convergence

Unobserved autocorrelated state variables 1

- Suppose state $x_{mt} = (x_{mt}^o, x_{mt}^u)$ where only x_{mt}^o is observed
- **Kasahara and Shimotsu (2009)** (for finite) and **Hu and Shum (2012)** (for continuous) give conditions for identification of transition probabilities $P(\cdot | x_{mt}^o, x_{mt}^u)$
- Given consistent $\hat{P}(\cdot | x_{mt}^o, x_{mt}^u)$ can apply 2-step estimator or NPL
 - Estimation is difficult

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- Computationally tractable estimation with unobserved state variables
- Two innovations:
 - Avoid matrix inversion in value function computation through finite dependence
 - Modified EM algorithm to integrate out distribution of unobserved states

EM algorithm 1

- “Expectation-Maximization”
 - Expectation: conditional probabilities of unobserved state given observables and parameters updated
 - Maximization: maximize likelihood as though unobserved state observed
 - Repeated until convergence
- Setup: observe x , missing s , complete $z = (x, s)$
- Joint likelihood $p(x, z|\theta)$
- Marginal likelihood $L(\theta; x) = p(x|\theta) = E[p(x, z|\theta)|x, \theta]$
- Difficulty:

$$E[p(x, s|\theta)|x, \theta] = \int_S p(x, s|\theta)p(s|x; \theta)ds$$

might be hard to compute

- Steps:

EM algorithm 2

- Initial θ^0
 - Expectation: calculate $p(s|x; \theta^0)$
 - Maximization: $\theta^1 = \arg \max_{\theta} \int_S p(x, s|\theta)p(s|x; \theta^0)ds$
 - Iterate to convergence
- Pros: stable – each iteration guaranteed to increase likelihood
 - Cons: slow?
 - Arcidiacono and Miller (2011):
 - Not slow when using finite dependence so that maximization step fast
 - Show 2-step version of EM algorithm possible, i.e. can estimate $p(s|x; \theta)$ without estimating θ

$$p(s|x; \theta) = \frac{p(x, s|\theta)}{\sum_{s'} p(x, s'|\theta)}$$

Can use empirical probabilities in place of model probabilities within EM algorithm

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