

# Estimating Production Functions

## Methodology Extensions

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# Mistakes have been made

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Akerberg,  
Caves, and  
Frazer (2015)

Gandhi,  
Navarro, and  
Rivers (2016)

References

Control function estimators of production functions have repeatedly been used without fully thinking through the underlying model and assumptions

- 1 Colinearity of flexible inputs with each other
  - Pointed out by [Akerberg, Caves, and Frazer \(2015\)](#)
- 2 Lack of relevant instrument for flexible input
  - Pointed out by [Gandhi, Navarro, and Rivers \(2016\)](#)
- 3 Heterogeneous markups are incompatible with the monotonicity assumption
  - Mistake in [De Loecker and Warzynski \(2012\)](#) (1337 citations), repeated in [De Loecker, Eeckhout, and Unger \(2020\)](#) (1268 citations)
  - Pointed out by [Doraszelski and Jaumandreu \(2019\)](#), [Doraszelski and Jaumandreu \(2021\)](#)

## Extensions

- **Levinsohn and Petrin (2003)**: investment often zero, so use other inputs instead of investment to form control function
- **Akerberg, Caves, and Frazer (2015)**: control function often collinear with  $l_{it}$  – for it not to be must be firm specific unobservables affecting  $l_{it}$  (but not investment / other input or else demand not invertible and cannot form control function)
- **Gandhi, Navarro, and Rivers (2016)**: relax scalar unobservable in investment / other input demand
- **Wooldridge (2009)**: more efficient joint estimation
- **Maican (2006) and Doraszelski and Jaumandreu (2013)**: endogenous productivity

## Section 1

# Ackerberg, Caves, and Frazer (2015)

# Akerberg, Caves, and Frazer (2015): contributions

## Akerberg, Caves, and Frazer (2015)

Collinearity in OP

ACF estimator

Relation to dynamic  
panel

Simulations

## Gandhi, Navarro, and Rivers (2016)

## References

- Document collinearity problem in OP and **Levinsohn and Petrin (2003)**
  - Need  $l_{it}, f_{it}(k_{it}, i_{it})$  not collinear, i.e. something causes variation in  $l$ , but not  $k$
- Propose alternative estimator
- Relates estimator to dynamic panel (**Blundell and Bond, 2000**) approach

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<sup>0\*</sup>These slides are based on the working paper version **Akerberg, Caves, and Frazer (2006)**.

## Collinearity in OP 1

- OP assume  $i_{it} = I_t(k_{it}, \omega_{it})$
- Symmetry, parsimony suggest  $l_{it} = L_t(k_{it}, \omega_{it})$
- Then  $l_{it} = L_t(k_{it}, I_t^{-1}(k_{it}, i_{it})) = g_t(k_{it}, i_{it})$

$$y_{it} = \beta_l l_{it} + f_t(k_{it}, i_{it}) + \epsilon_{it}$$

$l_{it}$  collinear with  $f_t(k_{it}, i_{it})$

- Worse in [Levinsohn and Petrin \(2003\)](#)
  - Uses other input  $m_{it}$  to form control function

$$y_{it} = \beta_l l_{it} + \beta_k k_{it} + \beta_m m_{it} + \omega_{it} + \epsilon_{it}$$

$$m_{it} = M_t(k_{it}, \omega_{it})$$

- Even less reason to treat labor demand differently than other input demand
- Collinearity still problem with parametric input demand

## Collinearity in OP 2

- Plausible models that do not solve collinearity
  - Input price data
    - Must include in control function to preserve scalar unobservable
    - Same logic above implies  $m$  and  $l$  are functions of both prices, so still collinear
  - Adjustest costs in labor
    - Need to add  $l_{it-1}$  to control function
  - Change in timing assumptions
  - Measurement error in  $l$  (but not  $m$ )
    - Solves collinearity, but makes  $\hat{\beta}_l$  inconsistent
- Potential model change that removes collinearity
  - Optimization error in  $l$  (but not  $m$ )
  - $m$  chosen,  $l$  specific shock revealed,  $l$  chosen
  - OP only:  $l_{it}$  chosen at  $t - 1/2$ ,  $l_{it} = L_t(\omega_{it-1/2}, k_{it})$ ,  $i_{it}$  chosen at  $t$

## ACF estimator

- Idea: like capital, labor is harder to adjust than other inputs
- Model:  $l_{it}$  chosen at time  $t - 1/2$ ,  $m_{it}$  at time  $t$ 
  - Implies  $m_t = M_t(k_{it}, l_{it}, \omega_{it})$

- Estimation:

$$\textcircled{1} \quad y_{it} = \underbrace{\beta_k k_{it} + \beta_l l_{it} + f_t(m_{it}, k_{it}, l_{it})}_{\equiv \Phi_t(m_{it}, k_{it}, l_{it})} + \epsilon_{it} \text{ gives}$$

$$\hat{\omega}_{it}(\beta_k, \beta_l) = \hat{\Phi}_{it} - \beta_k k_{it} - \beta_l l_{it}$$

- $\textcircled{2}$  Moments from timing and Markov process for  $\omega_{it}$  assumptions:

$$\omega_{it} = E[\omega_{it} | \omega_{it-1}] + \xi_{it}$$

- $E[\xi_{it} | k_{it}] = 0$  as in OP
- $E[\xi_{it} | l_{it-1}] = 0$  from new timing assumption
- $\hat{\xi}_{it}(\beta_k, \beta_l)$  as residual from nonparametric regression of  $\hat{\omega}_{it}$  on  $\hat{\omega}_{it-1}$
- Can add moments based on  $E[\epsilon_{it} | \mathcal{I}_{it}] = 0$



## Relation to dynamic panel estimators

- Both derive moment conditions from assumptions about timing and information set of firm
- Dealing with  $\omega$ 
  - Dynamic panel: AR(1) assumption allows quasi-differencing
  - Control function: makes  $\omega$  estimable function of observables
- Dynamic panel allows fixed effects, does not make assumptions about input demand
- Control function allows more flexible process for  $\omega_{it}$

# Simulations

- DGPS:
  - ① Consistent with their model, but not LP
  - ② Consistent with both
  - ③ Combination that consistent with neither
- Add measurement error to materials

# Simulation Results

TABLE I  
MONTE CARLO RESULTS<sup>a</sup>

Meas. Error	ACF				LP			
	$\beta_l$		$\beta_k$		$\beta_l$		$\beta_k$	
	Coef.	Std. Dev.	Coef.	Std. Dev.	Coef.	Std. Dev.	Coef.	Std. Dev.
	<i>DGP1—Serially Correlated Wages and Labor Set at Time <math>t - b</math></i>							
0.0	0.600	0.009	0.399	0.015	0.000	0.005	1.121	0.028
0.1	0.596	0.009	0.428	0.015	0.417	0.009	0.668	0.019
0.2	0.602	0.010	0.427	0.015	0.579	0.008	0.488	0.015
0.5	0.629	0.010	0.405	0.015	0.754	0.007	0.291	0.012
	<i>DGP2—Optimization Error in Labor</i>							
0.0	0.600	0.009	0.400	0.016	0.600	0.003	0.399	0.013
0.1	0.604	0.010	0.408	0.016	0.677	0.003	0.332	0.011
0.2	0.608	0.011	0.410	0.015	0.725	0.003	0.289	0.010
0.5	0.620	0.013	0.405	0.017	0.797	0.003	0.220	0.010
	<i>DGP3—Optimization Error in Labor and Serially Correlated Wages and Labor Set at Time <math>t - b</math> (DGP1 plus DGP2)</i>							
0.0	0.596	0.006	0.406	0.014	0.473	0.003	0.588	0.016
0.1	0.598	0.006	0.422	0.013	0.543	0.004	0.522	0.014
0.2	0.601	0.006	0.428	0.012	0.592	0.004	0.473	0.012
0.5	0.609	0.007	0.431	0.013	0.677	0.005	0.386	0.012

<sup>a</sup>1000 replications. True values of  $\beta_l$  and  $\beta_k$  are 0.6 and 0.4, respectively. Standard deviations reported are of parameter estimates across the 1000 replications.

## Section 2

# Gandhi, Navarro, and Rivers (2016)

# Gandhi, Navarro, and Rivers (2016)

- Show that control function method is not nonparametrically identified when there are flexible inputs
- Propose alternate estimate that uses data on input shares and information from firm's first order condition
- Show that value-added and gross output production functions are incompatible
- Application to Colombia and Chile

# Assumptions

- 1 Hicks neutral productivity  $Y_{jt} = e^{\omega_{jt} + \epsilon_{jt}} F_t(L_{jt}, K_{jt}, M_{jt})$
- 2  $\omega_{jt}$  Markov,  $\epsilon_{jt}$  i.i.d.
- 3  $K_{jt}$  and  $L_{jt}$  determined at  $t - 1$ ,  $M_{jt}$  determined flexibly at  $t$ 
  - $K$  and  $L$  play same role in the model, so after this slide I will drop  $L$
- 4  $M_{jt} = \mathbb{M}_t(L_{jt}, K_{jt}, \omega_{jt})$ , monotone in  $\omega_{jt}$

## Reduced form

- Let  $h(\omega_{jt-1}) = E[\omega_{jt} | \omega_{jt-1}]$ ,  $\eta_{jt} = \omega_{jt} - h(\omega_{jt-1})$
- log output

$$\begin{aligned} y_{jt} &= f_t(k_{jt}, m_{jt}) + \omega_{jt} + \epsilon_{jt} \\ &= f_t(k_{jt}, m_{jt}) + \underbrace{h(\mathbb{M}_{t-1}^{-1}(k_{jt-1}, m_{jt-1}))}_{=h_{t-1}(k_{jt-1}, m_{jt-1})} + \eta_{jt} + \epsilon_{jt} \end{aligned}$$

- Assumptions imply

$$E[\eta_{jt} | \underbrace{k_{jt}, k_{jt-1}, m_{jt-1}, \dots, k_{j1}, m_{j1}}_{=\Gamma_{jt}}] = 0$$

- Reduced form

$$E[y_{jt} | \Gamma_{jt}] = E[f_t(k_{jt}, m_{jt}) | \Gamma_{jt}] + h_{t-1}(k_{jt-1}, m_{jt-1}) \quad (1)$$

- Identification: given observed  $E[y_{jt} | \Gamma_{jt}]$  is there a unique  $f_t, h_{t-1}$  that satisfies (3)?

## Example: Cobb-Douglas 1

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Akerberg,  
Caves, and  
Frazer (2015)

Gandhi,  
Navarro, and  
Rivers (2016)

Identification  
problem

Identification from  
first order conditions

Value added vs gross  
production

Empirical results

References

- Let  $f_t(k, m) = \beta_k k + \beta_m m$
- Assume firm is takes prices as given
- First order condition for  $m$  gives

$$m = \text{constant} + \frac{\beta_k}{1 - \beta_m} k + \frac{1}{1 - \beta_m} \omega$$

- Put into reduced form

$$E[y_{jt} | \Gamma_{jt}] = C + \frac{\beta_k}{1 - \beta_m} k_{jt} + \frac{\beta_m}{1 - \beta_m} E[\omega_{jt} | \Gamma_{jt}] + h_{t-1}(k_{jt-1}, m_{jt-1}) \quad (2)$$

- $\omega$  Markov and  $\omega_{jt-1} = \mathbb{M}_{t-1}^{-1}(k_{jt-1}, m_{jt-1})$  implies

$$\begin{aligned} E[\omega_{jt} | \Gamma_{jt}] &= E[\omega_{jt} | \omega_{jt-1} = \mathbb{M}_{t-1}^{-1}(k_{jt-1}, m_{jt-1})] = \\ &= h_{t-1}(k_{jt-1}, m_{jt-1}) \end{aligned}$$



## Example: Cobb-Douglas 2

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Akerberg,  
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Frazer (2015)

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Navarro, and  
Rivers (2016)

Identification  
problem

Identification from  
first order conditions

Value added vs gross  
production

Empirical results

References

- Which leaves

$$E[y_{jt} | \Gamma_{jt}] = \text{constant} + \frac{\beta_k}{1 - \beta_m} k_{jt} + \frac{1}{1 - \beta_m} h_{t-1}(k_{jt-1}, m_{jt-1}) \quad (3)$$

from which  $\beta_k, \beta_m$  are not identified

- Rank condition fails,  $E[m_{jt} | \Gamma_{jt}]$  is colinear with  $h_{t-1}(k_{jt-1}, m_{jt-1})$
- After conditioning on  $k_{jt}, k_{jt-1}, m_{jt-1}$ , only variation in  $m_{jt}$  is from  $\eta_{jt}$ , but this is uncorrelated with the instruments

## Identification from first order conditions 1

- Since  $m$  flexible, it satisfies a simple static first order condition,

$$\rho_t = p_t \frac{\partial F_t}{\partial M} E[e^{\epsilon_{jt}}] e^{\omega_{jt}}$$

$$\log \rho_t = \log p_t + \log \frac{\partial F_t}{\partial M}(k_{jt}, m_{jt}) + \log E[e^{\epsilon_{jt}}] + \omega_{jt}$$

- Problem: prices often unobserved, endogenous  $\omega$
- Solution: difference from output equation to eliminate  $\omega$ , rearrange so that it involves only the value of materials and the value of output (which are often observed)

$$\underbrace{s_{jt}}_{\equiv \log \frac{\rho_t M_{jt}}{p_t Y_{jt}}} = \log \underbrace{G_t(k_{jt}, m_{jt})}_{\equiv \left( M_t \frac{\partial F_t}{\partial M} \right) / F_t} + \log \underbrace{E[e^{\epsilon_{jt}}]}_{\mathcal{E}} - \epsilon_{jt}$$

## Identification from first order conditions 2

- Identifies elasticity up to scale,  $G_t \mathcal{E}$  and  $\epsilon_{jt}$  which identify  $\mathcal{E}$
- Integrating,

$$\int_{m_0}^{m_{jt}} G_t(k_{jt}, m)/m = f_t(k_{jt}, m_{jt}) + c_t(k_{jt})$$

identifies  $f$  up to location

- Output equation

$$y_{jt} = \int_{m_0}^{m_{jt}} \tilde{G}_t(k_{jt}, m)/m - c_t(k_{jt}) + \omega_{jt} + \epsilon_{jt}$$

$$-c_t(k_{jt}) + \omega_{jt} = y_{jt} - \underbrace{\int_{m_0}^{m_{jt}} \tilde{G}_t(k_{jt}, m)/m - \epsilon_{jt}}_{\equiv \mathcal{Y}_{jt}}$$

# Identification from first order conditions 3

where the things on the right have already been identified

- Identify  $c_t$  from

$$\mathcal{Y}_{jt} = -c_t(k_{jt}) + \tilde{h}_t(\mathcal{Y}_{jt-1}, k_{jt-1}) + \eta_{jt}$$

# Value added vs gross production

- Value added:

$$\begin{aligned} VA_{jt} &= p_t Y_{jt} - \rho_t M_{jt} \\ &= p_t F_t(K_{jt}, M_t(K_{jt}, \omega_{jt})) e^{\omega_{jt} + \epsilon_{jt}} - \rho_t M_t(K_{jt}, \omega_{jt}) \end{aligned}$$

- Envelope theorem implies  
elasticity $_{e^\omega}^Y \approx$  elasticity $_{e^\omega}^{VA} (1 - \frac{\rho_t M_{jt}}{p_t Y_{jt}})$

## Problems

- Production Hicks-neutral productivity does not imply value-added Hicks-neutral productivity
- Ex-post shocks  $\epsilon_{jt}$  not accounted for in approximation

# Empirical results

- Look at tables
- Value-added estimates imply much more productivity dispersion than gross (90-10) ratio of 4 vs 2

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