

Akerberg,
Caves, and
Frazer (2015)

Collinearity in OP
ACF estimator
Relation to dynamic
panel
Simulations

Gandhi,
Navarro, and
Rivers (2016)

Identification
problem
Identification from
first order conditions
Value added vs gross
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Empirical results

References

Estimating Production Functions

Methodology Extensions

Paul Schrimpf

UBC
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Mistakes have been made

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Control function estimators of production functions have repeatedly been used without fully thinking through the underlying model and assumptions

- 1 Colinearity of flexible inputs with each other
 - Pointed out by [Ackerberg, Caves, and Frazer \(2015\)](#)
- 2 Lack of relevant instrument for flexible input
 - Pointed out by [Gandhi, Navarro, and Rivers \(2016\)](#)
- 3 Heterogeneous markups are incompatible with the monotonicity assumption
 - Mistake in [De Loecker and Warzynski \(2012\)](#) (1337 citations), repeated in [De Loecker, Eeckhout, and Unger \(2020\)](#) (1268 citations)
 - Pointed out by [Doraszelski and Jaumandreu \(2019\)](#), [Doraszelski and Jaumandreu \(2021\)](#)

Critiques and extensions

- **Levinsohn and Petrin (2003)**: investment often zero, so use other inputs instead of investment to form control function
- **Akerberg, Caves, and Frazer (2015)**: control function often collinear with l_{it} – for it not to be must be firm specific unobservables affecting l_{it} (but not investment / other input or else demand not invertible and cannot form control function)
- **Gandhi, Navarro, and Rivers (2016)**: relax scalar unobservable in investment / other input demand
- **Wooldridge (2009)**: more efficient joint estimation
- **Maican (2006) and Doraszelski and Jaumandreu (2013)**: endogenous productivity

Section 1

Ackerberg, Caves, and Frazer (2015)

Akerberg, Caves, and Frazer (2015): contributions

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- Document collinearity problem in OP and **Levinsohn and Petrin (2003)**
 - Need $l_{it}, f_{it}(k_{it}, i_{it})$ not collinear, i.e. something causes variation in l , but not k
- Propose alternative estimator
- Relates estimator to dynamic panel (**Blundell and Bond, 2000**) approach

^{0*}These slides are based on the working paper version **Akerberg, Caves, and Frazer (2006)**.

Collinearity in OP 1

- OP assume $i_{it} = I_t(k_{it}, \omega_{it})$
- Symmetry, parsimony suggest $l_{it} = L_t(k_{it}, \omega_{it})$
- Then $l_{it} = L_t(k_{it}, I_t^{-1}(k_{it}, i_{it})) = g_t(k_{it}, i_{it})$

$$y_{it} = \beta_l l_{it} + f_t(k_{it}, i_{it}) + \epsilon_{it}$$

l_{it} collinear with $f_t(k_{it}, i_{it})$

- Worse in [Levinsohn and Petrin \(2003\)](#)
 - Uses other input m_{it} to form control function

$$y_{it} = \beta_l l_{it} + \beta_k k_{it} + \beta_m m_{it} + \omega_{it} + \epsilon_{it}$$

$$m_{it} = M_t(k_{it}, \omega_{it})$$

- Even less reason to treat labor demand differently than other input demand
- Collinearity still problem with parametric input demand

Collinearity in OP 2

- Plausible models that do not solve collinearity
 - Input price data
 - Must include in control function to preserve scalar unobservable
 - Same logic above implies m and l are functions of both prices, so still collinear
 - Adjustest costs in labor
 - Need to add l_{it-1} to control function
 - Change in timing assumptions
 - Measurement error in l (but not m)
 - Solves collinearity, but makes $\hat{\beta}_l$ inconsistent
- Potential model change that removes collinearity
 - Optimization error in l (but not m)
 - m chosen, l specific shock revealed, l chosen
 - OP only: l_{it} chosen at $t - 1/2$, $l_{it} = L_t(\omega_{it-1/2}, k_{it})$, i_{it} chosen at t

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ACF estimator

- Idea: like capital, labor is harder to adjust than other inputs
- Model: l_{it} chosen at time $t - 1/2$, m_{it} at time t
 - Implies $m_t = M_t(k_{it}, l_{it}, \omega_{it})$

- Estimation:

$$\textcircled{1} \quad y_{it} = \underbrace{\beta_k k_{it} + \beta_l l_{it} + f_t(m_{it}, k_{it}, l_{it})}_{\equiv \Phi_t(m_{it}, k_{it}, l_{it})} + \epsilon_{it} \text{ gives}$$

$$\hat{\omega}_{it}(\beta_k, \beta_l) = \hat{\Phi}_{it} - \beta_k k_{it} - \beta_l l_{it}$$

- $\textcircled{2}$ Moments from timing and Markov process for ω_{it} assumptions:

$$\omega_{it} = E[\omega_{it} | \omega_{it-1}] + \xi_{it}$$

- $E[\xi_{it} | k_{it}] = 0$ as in OP
- $E[\xi_{it} | l_{it-1}] = 0$ from new timing assumption
- $\hat{\xi}_{it}(\beta_k, \beta_l)$ as residual from nonparametric regression of $\hat{\omega}_{it}$ on $\hat{\omega}_{it-1}$
- Can add moments based on $E[\epsilon_{it} | \mathcal{I}_{it}] = 0$

Relation to dynamic panel estimators

- Both derive moment conditions from assumptions about timing and information set of firm
- Dealing with ω
 - Dynamic panel: AR(1) assumption allows quasi-differencing
 - Control function: makes ω estimable function of observables
- Dynamic panel allows fixed effects, does not make assumptions about input demand
- Control function allows more flexible process for ω_{it}

Simulations

- DGPS:
 - 1 Consistent with their model, but not LP
 - 2 Consistent with both
 - 3 Combination that consistent with neither
- Add measurement error to materials

Simulation Results

TABLE I
MONTE CARLO RESULTS^a

| Meas. Error | ACF | | | | LP | | | |
|---|-----------|-----------|-----------|-----------|-----------|-----------|-----------|-----------|
| | β_l | | β_k | | β_l | | β_k | |
| | Coef. | Std. Dev. | Coef. | Std. Dev. | Coef. | Std. Dev. | Coef. | Std. Dev. |
| <i>DGP1—Serially Correlated Wages and Labor Set at Time $t - b$</i> | | | | | | | | |
| 0.0 | 0.600 | 0.009 | 0.399 | 0.015 | 0.000 | 0.005 | 1.121 | 0.028 |
| 0.1 | 0.596 | 0.009 | 0.428 | 0.015 | 0.417 | 0.009 | 0.668 | 0.019 |
| 0.2 | 0.602 | 0.010 | 0.427 | 0.015 | 0.579 | 0.008 | 0.488 | 0.015 |
| 0.5 | 0.629 | 0.010 | 0.405 | 0.015 | 0.754 | 0.007 | 0.291 | 0.012 |
| <i>DGP2—Optimization Error in Labor</i> | | | | | | | | |
| 0.0 | 0.600 | 0.009 | 0.400 | 0.016 | 0.600 | 0.003 | 0.399 | 0.013 |
| 0.1 | 0.604 | 0.010 | 0.408 | 0.016 | 0.677 | 0.003 | 0.332 | 0.011 |
| 0.2 | 0.608 | 0.011 | 0.410 | 0.015 | 0.725 | 0.003 | 0.289 | 0.010 |
| 0.5 | 0.620 | 0.013 | 0.405 | 0.017 | 0.797 | 0.003 | 0.220 | 0.010 |
| <i>DGP3—Optimization Error in Labor and Serially Correlated Wages and Labor Set at Time $t - b$ (DGP1 plus DGP2)</i> | | | | | | | | |
| 0.0 | 0.596 | 0.006 | 0.406 | 0.014 | 0.473 | 0.003 | 0.588 | 0.016 |
| 0.1 | 0.598 | 0.006 | 0.422 | 0.013 | 0.543 | 0.004 | 0.522 | 0.014 |
| 0.2 | 0.601 | 0.006 | 0.428 | 0.012 | 0.592 | 0.004 | 0.473 | 0.012 |
| 0.5 | 0.609 | 0.007 | 0.431 | 0.013 | 0.677 | 0.005 | 0.386 | 0.012 |

^a1000 replications. True values of β_l and β_k are 0.6 and 0.4, respectively. Standard deviations reported are of parameter estimates across the 1000 replications.

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Section 2

Gandhi, Navarro, and Rivers (2016)

Gandhi, Navarro, and Rivers (2016)

- Show that control function method is not nonparametrically identified when there are flexible inputs
- Propose alternate estimate that uses data on input shares and information from firm's first order condition
- Show that value-added and gross output production functions are incompatible
- Application to Colombia and Chile

Assumptions

- 1 Hicks neutral productivity $Y_{jt} = e^{\omega_{jt} + \epsilon_{jt}} F_t(L_{jt}, K_{jt}, M_{jt})$
- 2 ω_{jt} Markov, ϵ_{jt} i.i.d.
- 3 K_{jt} and L_{jt} determined at $t - 1$, M_{jt} determined flexibly at t
 - K and L play same role in the model, so after this slide I will drop L
- 4 $M_{jt} = \mathbb{M}_t(L_{jt}, K_{jt}, \omega_{jt})$, monotone in ω_{jt}

Reduced form

- Let $h(\omega_{jt-1}) = E[\omega_{jt} | \omega_{jt-1}]$, $\eta_{jt} = \omega_{jt} - h(\omega_{jt-1})$
- log output

$$\begin{aligned} y_{jt} &= f_t(k_{jt}, m_{jt}) + \omega_{jt} + \epsilon_{jt} \\ &= f_t(k_{jt}, m_{jt}) + \underbrace{h(\mathbb{M}_{t-1}^{-1}(k_{jt-1}, m_{jt-1}))}_{=h_{t-1}(k_{jt-1}, m_{jt-1})} + \eta_{jt} + \epsilon_{jt} \end{aligned}$$

- Assumptions imply

$$E[\eta_{jt} | \underbrace{k_{jt}, k_{jt-1}, m_{jt-1}, \dots, k_{j1}, m_{j1}}_{=\Gamma_{jt}}] = 0$$

- Reduced form

$$E[y_{jt} | \Gamma_{jt}] = E[f_t(k_{jt}, m_{jt}) | \Gamma_{jt}] + h_{t-1}(k_{jt-1}, m_{jt-1}) \quad (1)$$

- Identification: given observed $E[y_{jt} | \Gamma_{jt}]$ is there a unique f_t, h_{t-1} that satisfies (3)?

Example: Cobb-Douglas 1

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- Let $f_t(k, m) = \beta_k k + \beta_m m$
- Assume firm is takes prices as given
- First order condition for m gives

$$m = \text{constant} + \frac{\beta_k}{1 - \beta_m} k + \frac{1}{1 - \beta_m} \omega$$

- Put into reduced form

$$E[y_{jt} | \Gamma_{jt}] = C + \frac{\beta_k}{1 - \beta_m} k_{jt} + \frac{\beta_m}{1 - \beta_m} E[\omega_{jt} | \Gamma_{jt}] + h_{t-1}(k_{jt-1}, m_{jt-1}) \quad (2)$$

- ω Markov and $\omega_{jt-1} = \mathbb{M}_{t-1}^{-1}(k_{jt-1}, m_{jt-1})$ implies

$$\begin{aligned} E[\omega_{jt} | \Gamma_{jt}] &= E[\omega_{jt} | \omega_{jt-1} = \mathbb{M}_{t-1}^{-1}(k_{jt-1}, m_{jt-1})] = \\ &= h_{t-1}(k_{jt-1}, m_{jt-1}) \end{aligned}$$

Example: Cobb-Douglas 2

- Which leaves

$$E[y_{jt} | \Gamma_{jt}] = \text{constant} + \frac{\beta_k}{1 - \beta_m} k_{jt} + \frac{1}{1 - \beta_m} h_{t-1}(k_{jt-1}, m_{jt-1}) \quad (3)$$

from which β_k, β_m are not identified

- Rank condition fails, $E[m_{jt} | \Gamma_{jt}]$ is colinear with $h_{t-1}(k_{jt-1}, m_{jt-1})$
- After conditioning on $k_{jt}, k_{jt-1}, m_{jt-1}$, only variation in m_{jt} is from η_{jt} , but this is uncorrelated with the instruments

Identification from first order conditions 1

- Since m flexible, it satisfies a simple static first order condition,

$$\rho_t = p_t \frac{\partial F_t}{\partial M} E[e^{\epsilon_{jt}}] e^{\omega_{jt}}$$

$$\log \rho_t = \log p_t + \log \frac{\partial F_t}{\partial M}(k_{jt}, m_{jt}) + \log E[e^{\epsilon_{jt}}] + \omega_{jt}$$

- Problem: prices often unobserved, endogenous ω
- Solution: difference from output equation to eliminate ω , rearrange so that it involves only the value of materials and the value of output (which are often observed)

$$\underbrace{s_{jt}}_{\equiv \log \frac{\rho_t M_{jt}}{p_t Y_{jt}}} = \log \underbrace{G_t(k_{jt}, m_{jt})}_{\equiv \left(M_t \frac{\partial F_t}{\partial M} \right) / F_t} + \log \underbrace{E[e^{\epsilon_{jt}}]}_{\mathcal{E}} - \epsilon_{jt}$$

Identification from first order conditions 2

- Identifies elasticity up to scale, $G_t \mathcal{E}$ and ϵ_{jt} which identify \mathcal{E}
- Integrating,

$$\int_{m_0}^{m_{jt}} G_t(k_{jt}, m)/m = f_t(k_{jt}, m_{jt}) + c_t(k_{jt})$$

identifies f up to location

- Output equation

$$y_{jt} = \int_{m_0}^{m_{jt}} \tilde{G}_t(k_{jt}, m)/m - c_t(k_{jt}) + \omega_{jt} + \epsilon_{jt}$$

$$-c_t(k_{jt}) + \omega_{jt} = y_{jt} - \underbrace{\int_{m_0}^{m_{jt}} \tilde{G}_t(k_{jt}, m)/m - \epsilon_{jt}}_{\equiv \mathcal{Y}_{jt}}$$

Identification from first order conditions 3

where the things on the right have already been identified

- Identify c_t from

$$y_{jt} = -c_t(k_{jt}) + \tilde{h}_t(y_{jt-1}, k_{jt-1}) + \eta_{jt}$$

Value added vs gross production

- Value added:

$$\begin{aligned} VA_{jt} &= p_t Y_{jt} - \rho_t M_{jt} \\ &= p_t F_t(K_{jt}, M_t(K_{jt}, \omega_{jt})) e^{\omega_{jt} + \epsilon_{jt}} - \rho_t M_t(K_{jt}, \omega_{jt}) \end{aligned}$$

- Envelope theorem implies
elasticity $_{e^\omega}^Y \approx$ elasticity $_{e^\omega}^{VA} (1 - \frac{\rho_t M_{jt}}{p_t Y_{jt}})$

Problems

- Production Hicks-neutral productivity does not imply value-added Hicks-neutral productivity
- Ex-post shocks ϵ_{jt} not accounted for in approximation

Empirical results

- Look at tables
- Value-added estimates imply much more productivity dispersion than gross (90-10) ratio of 4 vs 2

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