

Models of insurance demand

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Economics 565

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- 1 Starc (2014)
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- 2 Saltzman (2019)
- 3 Tebaldi, Torgovitsky, and Yang (2023)
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- 4 Einav et al. (2013)

References

- **Reviews:** Einav, Finkelstein, and Levin (2010), Einav and Finkelstein (2011), Gaynor, Ho, and Town (2015) sections 6 & 7, Einav, Finkelstein, and Mahoney (2021), Handel and Ho (2021)
- **BLP models of insurance demand:** Bundorf, Levin, and Mahoney (2012), Starc (2014)
- **Expected utility models of insurance demand:** Cardon and Hendel (2001), Einav et al. (2013)
- **Behavioral:** Handel (2013), Barseghyan et al. (2013), Handel and Kolstad (2015)

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Section 1

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References

- Health insurance industry concentrated
- Mergers often blocked by antitrust
 - Aetna & Humana
 - Anthem & Cigna
- What are the sources and consequences of insurer market power?
- Medigap insurance
- Estimate model of demand and firm pricing
- Results
 - Low demand elasticity, strong brand preferences
 - Average cost pricing would decrease premiums by 17%

Medigap 1

- Medicare has high deductibles & copays
 - Part A (hospitalization) deductible \approx \$1000
 - Part B (outpatient) copays 20%, no maximum
- Medigap provides extra coverage
- Set of plans regulated (price [and branding] is only characteristic chosen by firms)
- Open-enrollment period (within 6 months of enrolling in Medicare) price only based on age, gender, state, & smoking
- Minimum Loss Ratio: at least 65% of premiums must be used to cover claims
- Taxes vary within consumer state based on insurer state
- Data:
 - NAIC: insurer premiums, quantities, claims
 - MCBS: individual demographics, whether have any Medigap (but not which insurer & plan), claims

TABLE 1 Medicare Supplement Plans

| | A | B | C | D | E | F | G | H | I | J | K | L |
|-------------------------------|----|----|-----|----|----|-----|-----|----|----|-----|-----|------|
| Part A coinsurance | X | X | X | X | X | X | X | X | X | X | X | X |
| Part B coinsurance | X | X | X | X | X | X | X | X | X | X | 0.5 | 0.75 |
| Blood | X | X | X | X | X | X | X | X | X | X | 0.5 | 0.75 |
| Hospice | | | | | | | | | | | 0.5 | 0.75 |
| Skilled nursing | | | X | X | X | X | X | X | X | X | 0.5 | 0.75 |
| Part A deductible | | X | X | X | X | X | X | X | X | X | 0.5 | 0.75 |
| Part B deductible | | | X | | | X | | | | X | | |
| Part B excess charges | | | | | | X | 0.8 | | X | X | | |
| Foreign travel emergency | | | X | X | X | X | X | X | X | X | | |
| At-home recovery | | | | | | | X | | X | X | | |
| Preventative care coinsurance | X | X | X | X | | | | | | | | |
| Preventative care | | | | | | X | | | | X | | |
| Market share | 4% | 3% | 12% | 4% | 2% | 49% | 8% | 1% | 1% | 15% | 1% | 1% |

Source: NAIC data. Percentages do not add to one because of rounding. The “X” denotes plan coverage. The numbers in the final two columns represent the fraction of cost covered.

TABLE 2 Subsample Demographics

| | Subsample |
|--|------------------------------|
| Income | \$36,803.60 (\$57,278.53) |
| Self-reported health | 2.59 (1.11) |
| % Medigap | 23.62% (42.48%) |
| Private insurance premium paid given purchase | \$1,702.13 (\$1,440.84) |

Source: MCBS individual-level data. Standard deviations in parentheses. Self-reported health is ranked on a 1–5 scale where 1 is excellent and 5 is poor. Sample is restricted to consumers under 72 years of age. Medigap coverage is defined as having self-purchased private insurance. Consumers who are eligible for VA benefits (TRICARE) or Medicaid are not included in the subsample.

TABLE 3 Firms and Market Structure

| | National Market Share | Share of Active Markets | Average Premium |
|----------------------------------|--------------------------|----------------------------|--------------------|
| UnitedHealth | 0.46 | 0.98 | 1534.82 |
| Mutual of Omaha | 0.24 | 0.95 | 1398.38 |
| Conseco | 0.09 | 0.90 | 1615.26 |
| American Financial | 0.04 | 0.78 | 1630.09 |
| HCHSC | 0.03 | 0.05 | 1815.55 |
| Genworth Financial | 0.02 | 0.88 | 1517.81 |
| State Farm | 0.02 | 0.59 | 2159.99 |
| American Republic Mutual | 0.02 | 0.53 | 1323.05 |
| Universal American Financial | 0.01 | 0.79 | 1771.63 |
| Guarantee Trust | 0.01 | 0.50 | 1756.02 |
| Physicians Mutual | 0.01 | 0.68 | 1596.92 |
| USAA | 0.01 | 0.90 | 1677.31 |
| American National Financial | 0.01 | 0.67 | 1247.75 |
| Atlantic American | 0.01 | 0.63 | 1531.27 |
| Thrivent Financial for Lutherans | 0.01 | 0.38 | 1629.46 |
| State Mutual Company | 0.01 | 0.16 | 703.04 |
| Humana | 0.01 | 0.67 | 1247.23 |
| Liberty National | 0.01 | 0.88 | 1736.36 |

Source: NAIC plan-level data. The first column is the percentage of all Medigap plans sold by the firm. The second column gives the percentage of markets in which the firm offers any policy, and the third column is the average list premium.

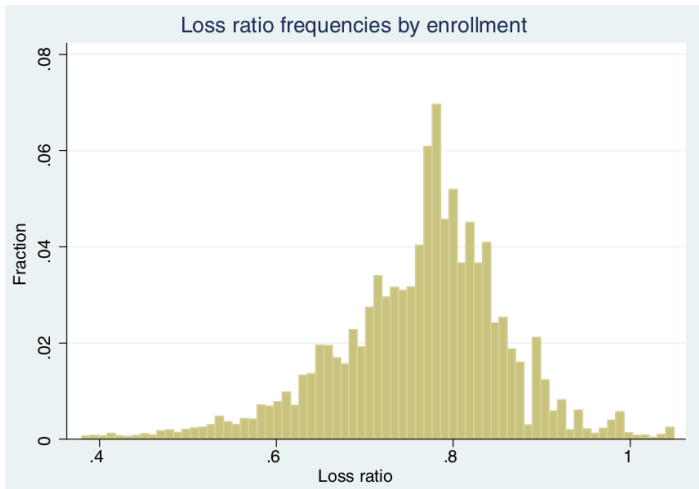
TABLE 4 Premiums and Claims by Plan

| Plan | % Load | Number of Policies | Weighted Premium (\$) | Weighted Claim (\$) | Enrollment |
|------|--------------------|--------------------|-----------------------|---------------------|----------------------|
| A | 0.2006 (0.8152) | 1403 | 1457.25 (743.26) | 1223.8 (1161.64) | 356.33 (1789.9) |
| B | 0.2192 (0.9803) | 1079 | 1562.31 (493.55) | 1218.03 (554.31) | 350.83 1159.28 |
| C | 0.2387 (0.5814) | 1764 | 1729.26 (389.37) | 1398.4 (460.22) | 908.62 (4973.19) |
| D | 0.3182 (0.5145) | 1822 | 1546.58 (459.44) | 1150.81 (451.93) | 325.11 (1057.07) |
| E | 0.3055 (0.3863) | 668 | 1691.22 (511.84) | 1235.19 (459.03) | 424.98 (1343.59) |
| F | 0.3213 (0.4834) | 3518 | 1518.81 (663.37) | 1170.77 (524.24) | 1908.30 (7807.93) |
| G | 0.3228 (0.4301) | 1936 | 1500.26 (446.44) | 1094.19 (380.99) | 591.17 (2034.61) |
| H | 0.2414 (0.4582) | 266 | 1379.37 (1379.37) | 1033.05 (493.11) | 394.96 (1326.39) |
| I | 0.3778 (0.3777) | 327 | 1675.13 (352.85) | 1252.45 (310.48) | 573.50 (1363.90) |
| J | 0.3539 (0.4335) | 716 | 1503.1 (380.23) | 1130.17 (341) | 2977.45 (9524.50) |
| K | 0.4543 (0.4739) | 308 | 712.59 (196.4) | 477.71 (183.35) | 176.87 (429.76) |
| L | 0.36 (0.5218) | 339 | 1183.35 (263.68) | 784.1 (784.3) | 251.01 (1502.79) |

Source: NAIC market-level data. Standard deviations in parentheses. Self-reported health is ranked on a 1–5 scale where 1 is excellent and 5 is poor. Column 2 describes the unweighted average load, and column 3 describes the number of policies. Columns 4 and 5 represent total enrollment-weighted averages of premiums and claims, respectively, with the weights reported in the final column.

FIGURE 1

LOSS RATIOS



Model 1

- Firm pricing:

$$\max_{p_{jfm}} \sum_j \left[\left(p_{jfm} - \underbrace{\gamma_{jfm}(\mathbf{p}_m)}_{\text{claims}} - \underbrace{a_{jfm}(\mathbf{p}_m)}_{\text{commissions}} \right) s_{jfm}(\mathbf{p}_m) M_m \right]$$

$$s.t. \gamma_{jgm}(\mathbf{p}_m) \geq 0.65 p_{jfm}$$

- Demand

- Consumer valuations:

$$v_{ijm} = x_j \beta_1 + b_f \beta_2 + x_m \beta_3 + \xi_{jfm} + \alpha p_{jfm} + \mu_{ijfm} + \epsilon_{ijfm}$$

- μ_{ijfm} = interactions between x_j and (z_i, ω_i)
- Claims:

$$\gamma_{ijfm} = \theta_0 + x_j \theta_1 + \underbrace{\omega_i \theta_2}_{\text{income}} + \underbrace{z_i \theta_3}_{\text{SRH}} + \epsilon_{ijm} + \eta_i$$

Estimation 1

- Demand estimation moments:
 - BLP market level data: $E[\zeta_{jfm} | \text{instruments}] = 0$
 - Retaliatory taxes
 - Average $p_{jff(-m)t}$
 - Expected claims given plan:

$$E[\gamma_{ifm} | J = j] = \theta_0 + x_j \theta_1 + E[\omega_i | J = j] \theta_2 E[z_i | J = j] \theta_3 + \varepsilon_{jm}$$

- Individual P(any Medigap), premium
- Pricing FOC used to estimate marginal costs (commissions)
 - Equality if MLR slack, inequality if binding or violated

TABLE 5 Demand Parameters

| | (1) | (2) | (3) | (4) |
|-----------------------------|---------------------|---------------------|---------------------|---------------------|
| Premium (in hundreds of \$) | -0.0767 (0.0070) | -0.1053 (0.0070) | -0.1049 (0.0070) | -0.0771 (0.0076) |
| B | 0.1007 (0.0819) | 0.1046 (0.0818) | 0.1043 (0.0818) | 0.0995 (0.0848) |
| C | 0.5484 (0.0864) | 0.5521 (0.0860) | 0.5520 (0.086) | 0.5434 (0.0900) |
| D | 0.5235 (0.0826) | 0.5350 (0.0826) | 0.5347 (0.0827) | 0.5247 (0.0845) |
| E | 0.4254 (0.1105) | 0.4395 (0.1107) | 0.4390 (0.1108) | 0.4315 (0.1129) |
| F | 1.5931 (0.0765) | 1.6028 (0.0764) | 1.6024 (0.0764) | 1.5987 (0.0782) |
| G | 0.8371 (0.0856) | 0.8478 (0.0856) | 0.8474 (0.0856) | 0.8471 (0.0873) |
| H | -0.4497 (0.1229) | -0.4403 (0.1232) | -0.4406 (0.1232) | -0.4455 (0.1309) |
| I | 0.1165 (0.1071) | 0.1236 (0.1072) | 0.1233 (0.1072) | 0.1219 (0.1119) |
| J | 1.8904 (0.0960) | 1.8999 (0.0961) | 1.8996 (0.0961) | 1.9010 (0.0985) |
| K | -1.4513 (0.1094) | -1.4460 (0.1096) | -1.4463 (0.1096) | -1.4439 (0.1161) |
| L | -1.0290 (0.1054) | -1.0224 (0.1057) | -1.0228 (0.1057) | -1.0206 (0.1114) |

Notes and Sources: MCBS data, NAIC data, and author calculations described in the text in detail. Brand dummies are included in the demand moments. Standard errors, adjusted for simulation error, are in parentheses.

TABLE 6 Additional Demand and Claim Parameters

| | (1) | (2) | (3) | (4) |
|---|---------------------|----------------------|----------------------|----------------------|
| Panel A. Parameters of claim equation | | | | |
| income | 0.0075 (0.0005) | 0.0012 (0.0008) | 0.0013 (0.0284) | 0.0082 (0.0011) |
| SRH | -0.0795 (0.0338) | 0.6696 (0.0324) | 0.6581 (1.2776) | 0.1643 (0.0122) |
| B | 0.3728 (0.0410) | 0.3234 (0.0410) | 0.3249 (0.0410) | 0.3170 (0.0480) |
| C | 0.5091 (0.0365) | 0.4750 (0.0354) | 0.4765 (0.0365) | 0.4857 (0.0357) |
| D | 0.2708 (0.0366) | 0.2426 (0.0357) | 0.2436 (0.0366) | 0.2793 (0.0361) |
| E | 0.2667 (0.0464) | 0.2574 (0.0457) | 0.2581 (0.0464) | 0.2670 (0.0451) |
| F | 0.2031 (0.0348) | 0.1980 (0.0338) | 0.1985 (0.0348) | 0.2367 (0.0340) |
| G | 0.2372 (0.036) | 0.2226 (0.0351) | 0.2233 (0.036) | 0.2422 (0.0353) |
| H | 0.2486 (0.064) | 0.2559 (0.0632) | 0.2561 (0.064) | 0.2045 (0.0612) |
| I | 0.1512 (0.0646) | 0.1550 (0.0640) | 0.1554 (0.0646) | 0.0955 (0.0637) |
| J | -0.0760 (0.0548) | -0.0443 (0.0535) | -0.0446 (0.0548) | -0.0521 (0.0523) |
| K | -0.9429 (0.0723) | -0.8876 (0.0715) | -0.8891 (0.0723) | -0.9487 (0.0673) |
| L | -0.3868 (0.0695) | -0.35419 (0.0638) | -0.35493 (0.0603) | -0.39877 (0.0625) |
| Panel B. Consumer demand heterogeneity | | | | |
| income*premium | 0.0002 (0.0000) | 0.0002 (0.0000) | 0.0002 (0.0000) | 0.0001 (0.1669) |
| Self-reported health*premium | | 0.0106 (0.0004) | 0.0104 (0.0055) | |
| Self-reported health*1(Medigap) | | | 0.0023 (0.1398) | |
| Self-reported health*1(United or Mutual of Omaha) | | | | 2.2397 (0.1669) |
| Panel C. Impact of estimates | | | | |
| mean elasticity | -1.1301 | -1.1227 | -1.1230 | -1.1338 |
| Mean derivative of claims w/r/t price | 0.0091 | 0.0716 | 0.0697 | 0.0940 |
| Value of AARP brand effect (in hundreds of \$) | 1.1863 | 1.1122 | 1.1131 | 2.0279 |

Source: NAIC market-level data, MCBS individual-level data, and author calculations described in the text in detail. Brand dummies are included in the demand side moments. Standard errors in parentheses. Self-reported health is ranked on a 1–5 scale where 1 is excellent and 5 is poor.

TABLE 7 Marginal Costs

| | Estimate | S.E. |
|-------------------------------------|----------|--------|
| Market Average, Unconstrained Model | 0.1942 | 0.0049 |
| Market Average | 0.1587 | 0.0011 |
| UnitedHealth | 0.0747 | 0.0000 |
| Mutual of Omaha | 0.1809 | 0.0252 |
| Conseco | 0.0814 | 0.0029 |
| American Financial | 0.1061 | 0.0010 |
| HCHSC | 0.0700 | 0.0002 |
| Genworth Financial | 0.1495 | 0.0003 |
| State Farm | 0.1630 | 0.0019 |
| American Republic Mutual | 0.1486 | 0.0001 |
| Universal American Financial | 0.1437 | 0.0004 |
| Guarantee Trust | 0.1506 | 0.0002 |
| Physicians Mutual | 0.1558 | 0.0002 |
| USAA | 0.1643 | 0.0002 |
| American National Financial | 0.1602 | 0.0001 |
| Atlantic American | 0.1624 | 0.0001 |
| Thrivent Financial for Lutherans | 0.1511 | 0.0002 |
| State Mutual Company | 0.2405 | 0.0000 |
| Humana | 0.1653 | 0.0001 |
| Liberty National | 0.1533 | 0.0006 |

Source: NAIC market-level data, MCBS individual-level data, and author calculations described in the text in detail. Standard errors are obtained using a bootstrap procedure that accounts for error in the demand estimates.

Consequences of market power

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TABLE 8 Average Cost and Claim Pricing

| | Average Claim | Average Cost |
|------------------------------------|---------------|--------------|
| Median premium (in hundreds of \$) | 8.8604 | 10.355 |
| Median % change in premium | -0.24241 | -0.17447 |
| Median % change in enrollment | 0.31231 | 0.21809 |
| Median compensating variation | 4.7237 | 3.5532 |
| Median CV net of profit loss | 2.0227 | 1.8458 |

Notes: The median premium paid is calculated as the median average premium paid across all state-year markets. The median percentage change in premium paid is calculated similarly. When noted, the change in total surplus includes both compensating variation and insurer profits. Compensating variation is calculated as the average across consumers within a market using the standard log-sum formula; the number reported is the median acrossmarkets.

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TABLE 9 Policy Counterfactuals

| | 65% MLR | 80% MLR | Mutual of Omaha Branding | United Branding |
|-------------------------------------|---------|---------|-----------------------------|--------------------|
| Median premium (in hundreds of \$) | 12.1970 | 11.4540 | 10.3520 | 10.2600 |
| Median % change in premium | -0.0573 | -0.1148 | -0.1999 | -0.2070 |
| Median % change in consumer surplus | 0.0320 | 0.0668 | -0.1823 | 0.0884 |

Notes: The median premium paid is calculated as the median average premium paid across all state-year markets. The median percentage change in premium paid is calculated similarly.

Source of market power

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TABLE A7 Fixed and Sunk Cost Estimates

| | Lower Bound | Upper Bound |
|--------------------------------|--|--|
| Sunk cost, UnitedHealth | \$99,261,645.01 (-\$1,530,902,861,706.31) | \$487,935,210.41 (-\$23,031,614,127.02) |
| Fixed cost, Mutual of Omaha | \$445,010.32 (-\$225,593.04) | \$796,342.56 (-\$3,578,033.82) |

TABLE A8 Marketing Expenditure and Advertising Value

| | United Health | Mutual of Omaha |
|---------------------------------------|---------------|-----------------|
| L.B. of sunk (fixed) cost/consumer | \$23.65 | \$8.37 |
| U.B. of sunk (fixed) cost/consumer | \$73.09 | \$14.81 |
| Average marginal cost/consumer | \$98.27 | \$238.67 |
| L.B. of total marketing cost/consumer | \$121.92 | \$247.05 |
| U.B. of total marketing cost/consumer | \$171.36 | \$253.48 |

Notes: Compensating variation is calculated as the average across consumers within a market using the standard log-sum formula; the number reported is the median across markets.

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Section 2

Saltzman (2019)

Saltzman (2019) “Demand for health insurance: Evidence from the California and Washington ACA exchanges”

- Estimate insurance demand
- Simulate impact of subsidies, mandate penalty, and mandate existence

ACA Exchanges

- Regulated state insurance exchanges
- Plan tiers based on expected percentage of health care costs covered
 - Bronze 60%, Silver 70%, Gold 80%, Platinum 90%
 - In California, plans standardized, elsewhere insurers can choose deductible, copay, etc
- Restrictions on price discrimination
 - Age: 64 year-old at most $3 \times$ 21 year old
 - Smoking: 50% more than non (prohibited in California)
 - Same price within geographic areas defined by states
- Mandatory to have some health insurance
 - Penalty: increased from $\max\{\$95, 1\% \text{income}\}$ to $\max\{\$625, 2.5\% \text{income}\}$ from 2014-2018, then \$0 after
 - Some exemptions
- Premium subsidies if income less than 400% of federal poverty level (price after subsidy is a max percentage of income ranging from 2%-9.5%)

Table 11
California exchange standard plan benefit designs (2014).

| | Bronze | Silver | Gold | Platinum | Silver 73 | Silver 87 | Silver 94 |
|---------------------|---------|---------|---------|----------|-----------|-----------|-----------|
| Actuarial value | 60% | 70% | 80% | 90% | 73% | 87% | 94% |
| Deductible | \$5,000 | \$2,000 | \$0 | \$0 | \$1,500 | \$500 | \$0 |
| Coinsurance | 30% | 20% | 20% | 10% | 20% | 15% | 10% |
| PCP copay | \$60 | \$45 | \$30 | \$20 | \$40 | \$15 | \$3 |
| Specialist copay | \$70 | \$65 | \$50 | \$40 | \$50 | \$20 | \$5 |
| Out-of-pocket limit | \$6,350 | \$6,350 | \$6,350 | \$4,000 | \$5,200 | \$2,250 | \$2,250 |

Notes: Table summarizes the standard plan benefit designs in the California exchange for the 2014 plan year. The silver 73, silver 87, and silver 94 plans are the enhanced versions of the basic silver plan and reduce cost sharing for consumers who qualify for cost sharing subsidies.

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$$U_{ij} = \alpha_i \underbrace{p_{ij}}_{\text{premium}} + \underbrace{x_j}_{\text{plan characteristics}} \underbrace{\beta}_{v_{ij}} + \underbrace{d_i}_{\text{household characteristics}} \varphi + \zeta_j + \epsilon_{ij}$$

$$\underbrace{U_{i0}}_{\text{uninsured}} = \alpha'_i \underbrace{\rho_i}_{\text{penalty}} + \epsilon_{i0}$$

- d_i includes d_{mi} = whether i subject to mandate, with coefficient φ_m
- Nested logit for ϵ_i with all plans in one nest, and uninsured in other (ϵ_{ij} are correlated with one another for $j \geq 1$ with correlation $\sqrt{1 - \lambda}$)

$$P(i \text{ chooses } j) = \frac{e^{V_{ij}/\lambda} \left(\sum_j e^{V_{ij}/\lambda} \right)^{\lambda-1}}{1 + \left(\sum_j e^{V_{ij}/\lambda} \right)^{\lambda-1}}$$

Table 2
Choice and demographic distribution by state.

| | California | | Washington | |
|----------------------------------|------------|-----------|------------|-----------|
| | Exchange | Uninsured | Exchange | Uninsured |
| Metals | | | | |
| Catastrophic | 0.7% | | 0.4% | |
| Bronze | 24.0% | | 36.6% | |
| Silver | 64.9% | | 55.1% | |
| Gold | 5.5% | | 7.7% | |
| Platinum | 4.8% | | 0.2% | |
| Network type | | | | |
| HMO | 45.7% | | 38.5% | |
| PPO | 45.1% | | 61.4% | |
| EPO | 9.2% | | 0.0% | |
| Access to free plan | 45.4% | 19.3% | 33.0% | 13.6% |
| Income | | | | |
| 0% to 138% of FPL | 2.9% | 2.8% | 5.0% | 4.3% |
| 138% to 150% of FPL | 15.0% | 5.4% | 8.5% | 4.6% |
| 150% to 200% of FPL | 33.8% | 20.5% | 30.3% | 18.0% |
| 200% to 250% of FPL | 17.4% | 16.2% | 18.7% | 17.3% |
| 250% to 400% of FPL | 22.7% | 29.6% | 25.0% | 30.9% |
| 400%+ of FPL | 8.2% | 25.4% | 12.5% | 25.0% |
| Subsidy eligibility | | | | |
| Premium tax credits | 90.7% | 74.6% | 85.5% | 75.0% |
| Cost sharing reduction subsidies | 68.5% | 44.9% | 61.4% | 44.2% |
| Penalty status | | | | |
| Exempt | 3.8% | 6.3% | 5.3% | 9.5% |
| Subject | 96.2% | 93.7% | 94.7% | 90.5% |
| Age | | | | |
| 0–17 | 4.8% | 3.2% | 0.3% | 2.9% |
| 18–25 | 10.4% | 20.9% | 8.5% | 19.1% |
| 26–34 | 15.7% | 25.5% | 17.5% | 25.2% |
| 35–44 | 15.6% | 17.0% | 17.4% | 19.9% |
| 45–54 | 24.4% | 17.8% | 22.6% | 16.6% |
| 55–64 | 29.0% | 15.4% | 33.8% | 16.3% |
| Gender | | | | |
| Female | 52.3% | 43.1% | 54.1% | 40.8% |
| Male | 47.7% | 56.9% | 45.9% | 59.2% |
| Race | | | | |
| Asian | | | 14.9% | 8.8% |
| Black/African American | | | 2.9% | 3.6% |
| Other Race | | | 5.4% | 12.1% |
| White | | | 76.8% | 75.5% |
| Smoking status | | | | |
| Non-smoker | | | 91.1% | 70.2% |
| Smoker | | | 8.9% | 29.8% |
| Year | | | | |
| 2014 | 48.9% | 58.9% | 48.0% | 56.5% |
| 2015 | 51.1% | 41.1% | 52.0% | 43.5% |

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Table 3

Insurers, plans, and premiums by state and year.

| | California | | Washington | |
|-----------------------------|------------|----------|------------|----------|
| | 2014 | 2015 | 2014 | 2015 |
| Insurers available | | | | |
| Minimum | 1.0 | 2.0 | 2.0 | 3.0 |
| Median | 5.0 | 5.0 | 6.0 | 7.0 |
| Average | 4.8 | 4.7 | 5.5 | 6.8 |
| Maximum | 6.0 | 6.0 | 7.0 | 8.0 |
| Plans available | | | | |
| Minimum | 5.0 | 10.0 | 16.0 | 21.0 |
| Median | 25.0 | 25.0 | 28.0 | 47.0 |
| Average | 24.6 | 24.5 | 26.2 | 45.8 |
| Maximum | 35.0 | 35.0 | 31.0 | 61.0 |
| Silver plan premiums | | | | |
| County average | \$309.70 | \$320.25 | \$306.00 | \$303.46 |
| Minimum | \$221.56 | \$230.31 | \$234.72 | \$218.55 |
| Maximum | \$480.59 | \$554.26 | \$369.11 | \$363.24 |
| Minimum second-lowest | \$253.27 | \$257.19 | \$260.01 | \$252.67 |
| Maximum second-lowest | \$422.58 | \$423.67 | \$312.61 | \$297.00 |

Notes: The first two panels provide summary statistics on the number of insurers and plans available to consumers. The third panel shows variation in silver plan premiums for a 40-year old nonsmoker.

Results

Table 4

Estimated mean own-premium elasticities and semi-elasticities.

| | California | | Washington | |
|-------------------|------------|-----------------|------------|-----------------|
| | Elasticity | Semi-elasticity | Elasticity | Semi-elasticity |
| Overall | -9.1 | -21.8 | -7.2 | -19.9 |
| Income (% of FPL) | | | | |
| 0-138 | -8.8 | -21.3 | -10.7 | -28.6 |
| 138-250 | -9.7 | -23.1 | -7.3 | -20.3 |
| 250-400 | -8.2 | -20.0 | -6.6 | -18.5 |
| 400+ | -7.8 | -19.1 | -5.3 | -15.3 |
| Gender | | | | |
| Female | -8.8 | -21.0 | -6.8 | -18.9 |
| Male | -9.5 | -22.6 | -7.6 | -20.9 |
| Age | | | | |
| 18-34 | -13.1 | -27.9 | -10.0 | -24.9 |
| 35-54 | -9.3 | -19.9 | -7.5 | -18.7 |
| 55+ | -5.6 | -12.0 | -4.9 | -12.4 |
| Smoking status | | | | |
| Smoker | | | -10.3 | -27.6 |
| Non-smoker | | | -6.6 | -18.3 |
| Race | | | | |
| Asian | | | -8.2 | -22.1 |
| Black | | | -11.5 | -30.3 |
| White | | | -6.8 | -18.7 |

Notes: Table shows mean own-premium elasticities and semi-elasticities by demographic group. A plan's own-premium elasticity indicates the percentage change in enrollment for a 1% increase in its premium and is computed using Eq. (9). A plan's

Table 5

Estimated mean elasticities and semi-elasticities for exchange coverage.

| | California | | Washington | |
|-------------------|------------|-----------------|------------|-----------------|
| | Elasticity | Semi-elasticity | Elasticity | Semi-elasticity |
| Overall | -1.2 | -3.3 | -1.1 | -3.7 |
| Income (% of FPL) | | | | |
| 0-138 | -1.2 | -3.3 | -1.6 | -5.4 |
| 138-250 | -1.3 | -3.5 | -1.2 | -4.0 |
| 250-400 | -1.1 | -3.1 | -1.1 | -3.7 |
| 400+ | -1.0 | -2.9 | -0.9 | -3.1 |
| Gender | | | | |
| Female | -1.1 | -3.2 | -1.0 | -3.5 |
| Male | -1.2 | -3.4 | -1.1 | -3.9 |
| Age | | | | |
| 18-34 | -1.6 | -4.1 | -1.4 | -4.4 |
| 35-54 | -1.1 | -2.9 | -1.0 | -3.3 |
| 55+ | -0.7 | -1.7 | -0.7 | -2.2 |
| Smoking status | | | | |
| Smoker | | | -1.5 | -4.6 |
| Non-smoker | | | -1.0 | -3.1 |
| Race | | | | |
| Asian | | | -1.2 | -3.9 |
| Black | | | -1.7 | -5.2 |
| White | | | -1.1 | -3.3 |

Notes: Table shows mean elasticities and semi-elasticities for exchange coverage by demographic group. The mean elasticity for exchange coverage indicates the percentage change in exchange enrollment if all exchange premiums increase by 1% and is computed using Eq. (11). The mean semi-elasticity for exchange coverage indicates the percentage change in exchange enrollment if the exchange premium for a given demographic group increases by 1%.

Results

Table 6

Estimated parameters of non-premium plan characteristics.

| | California | Washington |
|----------------------|----------------------|----------------------|
| Actuarial value (AV) | 4.125*** (0.240) | 3.591*** (0.159) |
| HMO | -0.275*** (0.016) | 1.009*** (0.085) |
| Deductible ratio | | -0.096*** (0.008) |
| Max. OOP ratio | | 0.010 (0.009) |

Notes: ***Significant at the 1% level. **Significant at the 5% level. *Significant at the 10% level. Table shows parameter estimates for the non-premium plan characteristics, including the actuarial value, whether the plan is an HMO, the ratio of the plan's deductible to the maximum deductible in the plan's metal tier, and the ratio of the plan's out-of-pocket limit to the maximum out-of-pocket limit in the plan's metal tier. Parameters for the latter two variables cannot be estimated for California because of plan standardization. Robust standard errors that correct for potential misspecification are shown in parentheses (see p. 503 of [Wooldridge \(2010\)](#)).

Results

Table 10
Impact of repealing the individual mandate.

| | Percent change in exchange enrollment | | Percent change in consumer surplus | |
|----------------------|--|----------|---------------------------------------|----------|
| | ACA subsidies | Vouchers | ACA subsidies | Vouchers |
| California | | | | |
| 5% premium increase | -18.6% | -20.5% | 1.6% | -2.8% |
| 10% premium increase | -18.9% | -22.8% | 1.5% | -7.4% |
| 25% premium increase | -19.7% | -29.3% | 1.2% | -20.1% |
| Washington | | | | |
| 5% premium increase | -13.4% | -17.2% | 6.8% | -3.5% |
| 10% premium increase | -14.3% | -21.9% | 6.0% | -14.0% |
| 25% premium increase | -16.1% | -35.7% | 5.0% | -40.3% |

Notes: Table shows the impact on enrollment and average annual consumer surplus of repealing the individual mandate under a voucher subsidy and under ACA subsidies. Three alternative supply response scenarios are considered: a 5% premium increase, a 10% premium increase, and a 25% premium increase.

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Section 3

Tebaldi, Torgovitsky, and Yang (2023)

Tebaldi, Torgovitsky, and Yang (2023) “Nonparametric estimates of demand in the California health insurance exchanges”

- How much do logit / mixed logit assumptions influence demand estimates?
- Setting: California ACA exchange
- Nonparametric partially identified demand estimates

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References

- 19 rating regions (premiums vary across regions and are constant within)
- 4 tiers of insurance coverage
- Region, tier, & age specific premium = insurer chosen region, tier premium \times federal age adjustment
- Premium subsidies and cost-sharing reductions for low income individuals
- Mandated participation with tax penalty (penalty repealed in 2017)

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Panel (a): Characteristics by metal tier before cost-sharing reductions

| Tier | Annual deductible | Annual max out-of-pocket | Primary visit | E.R. visit | Specialist visit | Preferred drugs | Advertised AV ^(*) |
|----------|-------------------|--------------------------|---------------|------------|------------------|-----------------|------------------------------|
| Bronze | \$5,000 | \$6,250 | \$60 | \$300 | \$70 | \$50 | 60% |
| Silver | \$2,250 | \$6,250 | \$45 | \$250 | \$65 | \$50 | 70% |
| Gold | \$0 | \$6,250 | \$30 | \$250 | \$50 | \$50 | 79% |
| Platinum | \$0 | \$4,000 | \$20 | \$150 | \$40 | \$15 | 90% |

Panel (b): Silver plan characteristics after cost-sharing reductions

| Income (%FPL) | Annual deductible | Annual max out-of-pocket | Primary visit | E.R. visit | Specialist visit | Preferred drugs | Advertised AV ^(*) |
|---------------|-------------------|--------------------------|---------------|------------|------------------|-----------------|------------------------------|
| 200-250% FPL | \$1,850 | \$5,200 | \$40 | \$250 | \$50 | \$35 | 74% |
| 150-200% FPL | \$550 | \$2,250 | \$15 | \$75 | \$20 | \$15 | 88% |
| 100-150% FPL | \$0 | \$2,250 | \$3 | \$25 | \$5 | \$5 | 95% |

Source: <http://www.coveredca.com/PDFs/2015-Health-Benefits-Table.pdf> .

(*) : Actuarial value (AV) is advertised to consumers as a percentage of medical expenses covered by the plan.

Premium Variation

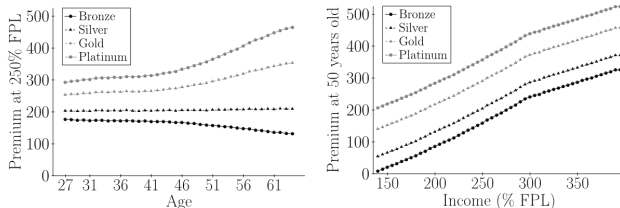


FIGURE 2.—Post-subsidy premium variation by age and income. *Notes:* Post-subsidy premiums shown are the median across insurers for rating region 16 (part of Los Angeles).

Model

- Individual i , plans $j \in \{0, 1, \dots, J\}$
- Valuations V_{ij} with premiums P_{ij} , with utility additively separable in premium

$$\max_j V_{ij} - P_{ij}$$

- Want to recover $f(V_{ij}|P_{ij}, M_i, X_i)$

Price Variation

- Premiums depend on market, M_i , and individual characteristics, X_i (age & income)
- Price variation within market will be used in estimation
- Price variation within market not present in typical demand estimation
- Appendix discusses modifications to use when there is not within market price variation

Target Parameters

- Density of valuation given observables $f(v|p, m, x)$
- Target parameters are functionals of this density, $\theta : \mathcal{F} \rightarrow \mathbb{R}^{d_\theta}$, e.g.

- Fraction that choose plan j if premiums were p^*

$$P(j|p^*, m, x) = \int \mathbf{1}\{v_j - p_j^* \geq v_k - p_k^* \forall k\} f(v|m, x) dv$$

- Change in consumer surplus from changing p to p^*

$$\Delta CS(p^*|m, x) = \int \max_j (v_j - p_j^*) f(v|m, x) dv - \int \max_j (v_j - p_j) f(v|m, x) dv$$

Assumptions

- W_i, Z_i subvectors of M_i, X_i
 - In application W_i is M_i and course age and income bins, Z_i is variation in age and income within bins
- Z_i is instrument
 - Exogenous:

$$f_{V|W,Z}(v|w, z) = f_{V|W,Z}(v|w, z') \quad (1)$$

- No relevance or rank assumption required, but size of identified will depend on instrument variation and relevance
- Support restrictions

$$\int_{\mathcal{V}^*(w)} f_{V|W,Z}(v|w, z) dz = 1 \quad (2)$$

e.g. at same prices, consumers prefer higher tier plan

$$\mathcal{V}^*(w) = \{v : v_4 \geq v_1\}$$

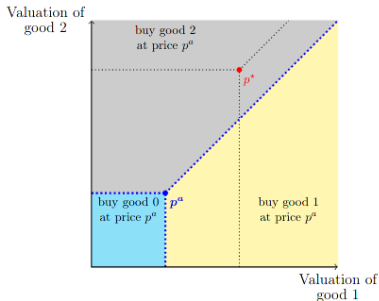
Identified Set

- Define $\mathcal{V}_j(p) = \{v : v_j - p_j \geq v_k - p_k \forall k\}$
- Observed shares = model shares:

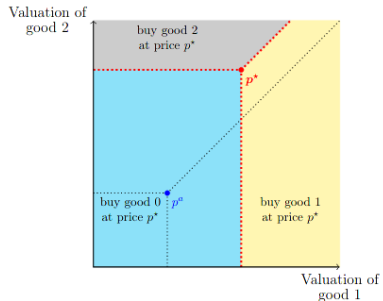
$$s_j(p, m, x; f) = \int_{\mathcal{V}_j(p)} f(v|p, m, x) dv \quad (3)$$

- Identified set $\mathcal{F}^* \equiv \{f \in \mathcal{F} : 1, 2, 3\}$
- Identified set for target parameter $\Theta^* \equiv \{\theta(f) : f \in \mathcal{F}^*\}$
- Goal : characterize and then estimate Θ^*

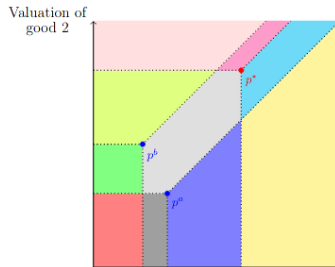
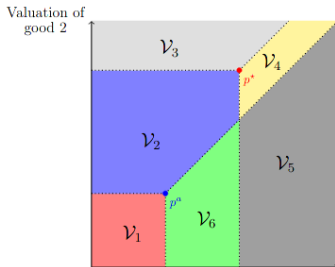
Figure 1: Partitioning the Space of Valuations



(a) Choices if prices were p^a .



(b) Choices if prices were p^* .



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- Observe p^a , counterfactual p^* , want $\theta(f) =$ share of good 2 at p^*
- Partition support of v into minimal relevant partition (c)
- We observe

$$s_0(m, p^a) = \int_{\mathcal{V}_1} f(v|m, p^a) dv$$

$$s_1(m, p^a) = \int_{\mathcal{V}_5 \cup \mathcal{V}_6} f(v|m, p^a) dv$$

$$s_2(m, p^a) = \int_{\mathcal{V}_2 \cup \mathcal{V}_3 \cup \mathcal{V}_4} f(v|m, p^a) dv$$

- Assume p exogenous, so $f(v|m, p^a) = f(v|m, p^*) = f(v|m)$ (i.e. $Z = p$)
- Let $\phi_\ell = \int_{\mathcal{V}_\ell} f(v|m) dv$, note that $s_2(m, p^*) = \phi_3$ is the parameter of interest
- Upper bound: $\max_\phi \phi_3$ s.t. observed shares

$$t_{ub}^* = \max_\phi \phi_3 \text{ s.t.}$$

$$\phi_1 = s_0(m, p^a)$$

$$\phi_2 + \phi_3 + \phi_4 = s_2(m, p^a)$$

$$\phi_5 + \phi_6 = s_1(m, p^a)$$

$$\phi_\ell \geq \forall \ell$$

- $t_{lb}^* = \min \phi_3$ gives lower bound, paper shows $[t_{lb}^*, t_{ub}^*]$ is the identified set

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Estimation 1

- Notation:

- \mathbb{V} is minimal relevant partition
- $\mathbb{V}_j(p)$ is subset of \mathbb{V} that rationalizes choice j given prices p
- $\phi(\mathcal{V}|m, x) = \int_{\mathcal{V}} f(v|m, x) dv$
- $\phi(\mathcal{V}|w, z) = \int_{\mathcal{V}} f(v|w, z) dv$

Estimation 1

- Just replace unknown population shares with observed market shares:

$$\min_{\phi \geq 0} \bar{\theta}(\phi) \text{ s.t.}$$

$$\hat{s}_j(m, x) = \sum_{\mathcal{V} \in \mathbb{V}_j(p(m, x))} \phi(\mathcal{V} | m, x) \quad \forall j$$

$$\phi_{\mathcal{V} | w, z}(\mathcal{V} | w, z) = \phi_{\mathcal{V} | w, z'}(\mathcal{V} | w, z') \quad \forall z, z', w, \mathcal{V}$$

$$\sum_{\mathcal{V} \in \mathbb{V}^*(w)} \phi_{\mathcal{V} | w, z}(\mathcal{V} | w, z) = 1 \quad \forall w, z$$

but might have no solution²

- Define:

$$\hat{Q}(\phi) = \sum_{j, m, x} \hat{\mathbb{P}}(m, x) \left| \hat{s}_j(m, x) - \sum_{\mathcal{V} \in \mathbb{V}_j(p(m, x))} \phi(\mathcal{V} | m, x) \right|$$

$$\text{and } \hat{Q}^* = \min_{\phi} \hat{Q}(\phi)$$

Estimation 2

- Relax problem to

$$\hat{t}_{lb}^* = \min_{\phi \geq 0} \bar{\theta}(\phi) \text{ s.t.}$$

$$\hat{Q}(\phi) \leq \hat{Q}^* + \eta$$

$$\phi_{\mathcal{V}|wz}(\mathcal{V}|w, z) = \phi_{\mathcal{V}|wz}(\mathcal{V}|w, z') \quad \forall z, z', w, \mathcal{V}$$

$$\sum_{\mathcal{V} \in \mathcal{V}^*(w)} \phi_{\mathcal{V}|wz}(\mathcal{V}|w, z) = 1 \quad \forall w, z$$

- Inference based on [Deb et al. \(2021\)](#)

²I think this is the reason, but the paper says “The purpose of this tuning parameter is to smooth out potential discontinuities caused by set convergence.”

Identifying Assumptions

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References

- California ACA pricing
 - 19 rating regions (premiums vary across regions and are constant within)
 - 4 tiers of insurance coverage
 - Region, tier, & age specific premium = insurer chosen region, tier premium \times federal age adjustment
 - Premium subsidies and cost-sharing reductions for low income individuals
- So price variation within a region due to age and income should be exogenous to demand shocks
- Assume that preferences for insurance do not depend on age or income within “coarse bins” (defined by 5 years and 50 percentage points of FPL)
- Support restriction: at equal prices, consumers prefer plan with more coverage

SUMMARY STATISTICS.

Panel (a): Distribution of bin characteristics

| | Mean | St. Dev. | P-10 | Median | P-90 |
|-------------------------|---------|----------|-------|--------|-------|
| Number of buyers | 85.32 | 91.07 | 14 | 55 | 195 |
| Age | 43.415 | 10.694 | 29 | 43 | 59 |
| Income (FPL%) | 243.991 | 72.037 | 155 | 230 | 355 |
| Takeup rate | 0.280 | 0.209 | 0.053 | 0.234 | 0.576 |
| Average premium paid | 174.495 | 89.324 | 68 | 162 | 298 |
| Share choosing Bronze | 0.065 | 0.073 | 0 | 0 | 0 |
| Share choosing Silver | 0.188 | 0.173 | 0 | 0 | 0 |
| Share choosing Gold | 0.015 | 0.021 | 0 | 0 | 0 |
| Share choosing Platinum | 0.012 | 0.018 | 0 | 0 | 0 |

Panel (b): Premiums and choice shares by age and income

| | Bronze | | Silver | | Gold | | Platinum | |
|--------------------------|---------|-------|---------|-------|---------|-------|----------|-------|
| | Premium | Share | Premium | Share | Premium | Share | Premium | Share |
| By age: | | | | | | | | |
| 27-34 | 120 | 0.050 | 174 | 0.122 | 229 | 0.010 | 272 | 0.010 |
| 35-49 | 117 | 0.058 | 181 | 0.175 | 248 | 0.013 | 299 | 0.011 |
| 50-64 | 104 | 0.086 | 207 | 0.259 | 321 | 0.022 | 409 | 0.016 |
| By income (FPL%): | | | | | | | | |
| 140-150 | 5 | 0.011 | 57 | 0.336 | 133 | 0.005 | 191 | 0.006 |
| 150-200 | 28 | 0.046 | 94 | 0.318 | 170 | 0.008 | 229 | 0.009 |
| 200-250 | 86 | 0.084 | 162 | 0.193 | 241 | 0.018 | 302 | 0.015 |
| 250-400 | 196 | 0.074 | 276 | 0.084 | 357 | 0.019 | 419 | 0.014 |

Note: Panel (a) reports statistic taken across the 30,007 bins in our main estimation sample. All statistics are weighted by number of potential buyers. For income, standard deviation means the standard deviation of the within-bin medians of income and average premium paid. In panel (b), premium is the average premium paid for buyers of a given age/income group.

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TABLE III
NONPARAMETRIC BOUNDS ON CHANGES IN CHOICE SHARES.

| \$10/month premium increase for | Change in probability of choosing | | | | | | | | | |
|------------------------------------|---|--------|--------|--------|--------|--------|--------|--------|----------|--------|
| | Any plan | | Bronze | | Silver | | Gold | | Platinum | |
| | LB | UB | LB | UB | LB | UB | LB | UB | LB | UB |
| | Panel (a): Full sample (140–400% FPL) | | | | | | | | | |
| All plans | -0.067 | -0.018 | -0.012 | -0.004 | -0.051 | -0.011 | -0.004 | -0.001 | -0.003 | -0.001 |
| Bronze | -0.011 | -0.002 | -0.047 | -0.009 | +0.004 | +0.044 | +0.000 | +0.028 | +0.000 | +0.023 |
| Silver | -0.050 | -0.003 | +0.001 | +0.124 | -0.165 | -0.017 | +0.001 | +0.121 | +0.000 | +0.097 |
| Gold | -0.003 | -0.000 | +0.000 | +0.005 | +0.000 | +0.010 | -0.013 | -0.003 | +0.000 | +0.011 |
| Platinum | -0.002 | -0.000 | +0.000 | +0.003 | +0.000 | +0.006 | +0.001 | +0.009 | -0.010 | -0.002 |
| | Panel (b): Lower income (140–250% FPL) | | | | | | | | | |
| All plans | -0.091 | -0.020 | -0.011 | -0.003 | -0.077 | -0.015 | -0.003 | -0.001 | -0.003 | -0.001 |
| Bronze | -0.009 | -0.001 | -0.046 | -0.008 | +0.004 | +0.044 | +0.000 | +0.027 | +0.000 | +0.023 |
| Silver | -0.076 | -0.005 | +0.001 | +0.178 | -0.237 | -0.021 | +0.001 | +0.173 | +0.000 | +0.141 |
| Gold | -0.002 | -0.000 | +0.000 | +0.004 | +0.000 | +0.009 | -0.011 | -0.002 | +0.000 | +0.010 |
| Platinum | -0.002 | -0.000 | +0.000 | +0.004 | +0.000 | +0.006 | +0.001 | +0.010 | -0.010 | -0.002 |
| | Panel (c): Higher income (250–400% FPL) | | | | | | | | | |
| All plans | -0.037 | -0.016 | -0.015 | -0.006 | -0.018 | -0.007 | -0.004 | -0.001 | -0.003 | -0.001 |
| Bronze | -0.013 | -0.003 | -0.049 | -0.009 | +0.003 | +0.045 | +0.000 | +0.029 | +0.000 | +0.023 |
| Silver | -0.016 | -0.001 | +0.001 | +0.053 | -0.072 | -0.012 | +0.001 | +0.054 | +0.000 | +0.040 |
| Gold | -0.003 | -0.000 | +0.000 | +0.006 | +0.000 | +0.012 | -0.016 | -0.004 | +0.000 | +0.013 |
| Platinum | -0.002 | -0.000 | +0.000 | +0.003 | +0.000 | +0.005 | +0.001 | +0.009 | -0.010 | -0.003 |

Note: Each pair of columns contains the estimated lower and upper bound for the change in choice probability of the choice indicated in columns in response to a \$10/month premium increase for the plan(s) indicated in the rows. The column "Any plan" means any choice $j \neq 0$, and the row "All plans" means all choices $j \neq 0$.

TABLE IV
SENSITIVITY TO ASSUMPTION IV.

| κ | Change in probability of purchasing coverage if all per-person premiums increase by \$10/month | | | | | |
|-----------|--|---------|---|---------|--|---------|
| | $\kappa_{\text{age}} = \kappa, \kappa_{\text{inc}} = 0$ | | $\kappa_{\text{age}} = 0, \kappa_{\text{inc}} = \kappa$ | | $\kappa_{\text{age}} = \kappa_{\text{inc}} = \kappa$ | |
| | LB | UB | LB | UB | LB | UB |
| 0 | -0.0674 | -0.0183 | -0.0674 | -0.0183 | -0.0674 | -0.0183 |
| 0.2 | -0.0691 | -0.0192 | -0.1076 | -0.0344 | -0.1017 | -0.0223 |
| 0.3 | -0.0699 | -0.0196 | -0.1227 | -0.0395 | -0.1083 | -0.0258 |
| 0.4 | -0.0705 | -0.0198 | -0.1355 | -0.0436 | -0.1191 | -0.0314 |
| 0.6 | -0.0718 | -0.0204 | -0.1556 | -0.0485 | -0.1415 | -0.0311 |
| $+\infty$ | -0.0865 | -0.0158 | -0.2602 | -0.0293 | -0.2798 | -0.0000 |

Note: Each pair of columns shows estimated lower and upper bounds on the change in choosing any inside choice ($j \neq 0$). The first pair adjusts κ_{age} , while keeping $\kappa_{\text{inc}} = 0$. The second pair adjusts κ_{inc} , while keeping $\kappa_{\text{age}} = 0$. The third pair adjusts both κ_{inc} and κ_{age} simultaneously.

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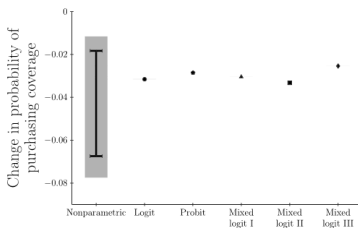
Saltzman (2019)

Tebaldi, Torgovitsky, and Yang (2023)

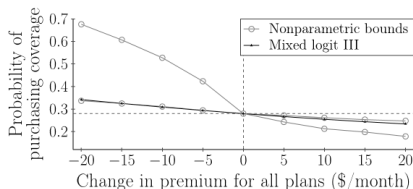
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Einav et al. (2013)

References



(a) Change in participation in response to a \$10 decrease in subsidies



(b) Change in participation in response to different sized changes in subsidies

FIGURE 3.—Comparison to parametric models. *Notes:* Top panel: Bound and point estimates are shown in solid black, and 95% confidence intervals are indicated with grey shading. The confidence interval for the logit and probit models are too narrow to be visible. Bottom panel: Nonparametric upper and lower bounds on the overall probability of purchasing coverage (choosing $j \neq 0$) for each price change are shown with light grey circles. Corresponding point estimates from mixed logit III are shown in black triangles.

TABLE V
ELASTICITIES.

| 1% premium increase for | | % change in probability of choosing | | | | | | |
|-------------------------|--|-------------------------------------|--------|-----------------------|---------|---------|-----------------------|--|
| | | Outside | | High deductible | | | Low deductible | |
| | | Bounds/Point estimate | | Bounds/Point estimate | | | Bounds/Point estimate | |
| High deductible | Nonparametric | +0.025 | +0.169 | -9.797 | -1.707 | +0.256 | +2.710 | |
| | $\kappa_{age} = 0.4$ $\kappa_{inc} = 0$ | +0.023 | +0.190 | -10.438 | -1.787 | +0.273 | +3.002 | |
| | $\kappa_{age} = \infty$ $\kappa_{inc} = 0$ | +0.000 | +0.282 | -11.369 | -1.051 | +0.074 | +3.452 | |
| | $\kappa_{age} = 0$ $\kappa_{inc} = 0.4$ | +0.073 | +0.387 | -10.046 | -2.632 | +0.152 | +2.707 | |
| | $\kappa_{age} = 0$ $\kappa_{inc} = \infty$ | +0.112 | +0.898 | -10.646 | -2.292 | +0.077 | +2.727 | |
| | Logit | +0.154 | | -1.997 | | +0.154 | | |
| | Probit | +0.152 | | -1.902 | | +0.200 | | |
| | Mixed Logit I | +0.152 | | -1.966 | | +0.203 | | |
| | Mixed Logit II | +0.206 | | -4.411 | | +0.997 | | |
| | Mixed Logit III | +0.176 | | -4.039 | | +1.282 | | |
| Low deductible | Nonparametric | +0.207 | +1.530 | +1.364 | +54.251 | -15.491 | -1.956 | |
| | $\kappa_{age} = 0.4$ $\kappa_{inc} = 0$ | +0.197 | +1.583 | +1.922 | +59.219 | -16.178 | -2.183 | |
| | $\kappa_{age} = \infty$ $\kappa_{inc} = 0$ | +0.052 | +1.909 | +0.235 | +67.867 | -18.444 | -1.358 | |
| | $\kappa_{age} = 0$ $\kappa_{inc} = 0.4$ | +0.472 | +3.064 | +0.955 | +52.746 | -17.351 | -3.638 | |
| | $\kappa_{age} = 0$ $\kappa_{inc} = \infty$ | +0.449 | +5.851 | +0.195 | +63.914 | -20.367 | -2.288 | |
| | Logit | +0.641 | | +0.641 | | -3.549 | | |
| | Probit | +0.544 | | +1.200 | | -2.455 | | |
| | Mixed Logit I | +0.619 | | +0.799 | | -3.426 | | |
| | Mixed Logit II | +0.281 | | +2.876 | | -3.135 | | |
| | Mixed Logit III | +0.182 | | +3.263 | | -3.187 | | |

Note: High deductible is Bronze and low deductible is a bundle consisting of Silver, Gold, and Platinum. See Appendix S6 for further details on implementation and computation.

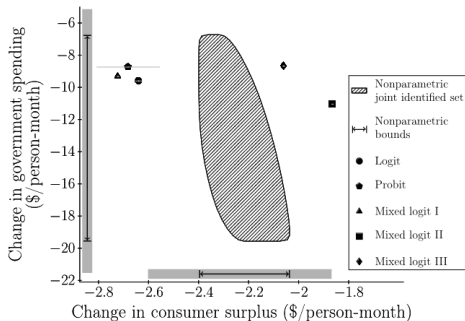


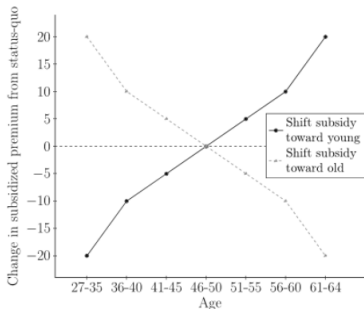
FIGURE 4.—Changes in consumer surplus and government spending from a \$10 decrease in subsidies. *Notes:* Bound and point estimates are shown in solid black. One-dimensional 95% confidence intervals are shown in grey vertical and horizontal bars.

TABLE VI

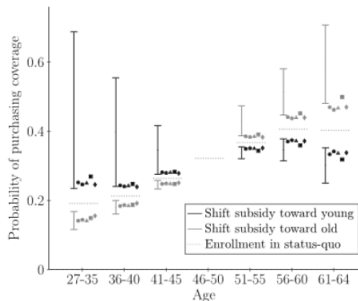
AGGREGATE IMPACTS FROM REDUCING PREMIUM SUBSIDIES BY \$10 PER MONTH.

| | 140–400% FPL | | 140–400% FPL | | 140–250% FPL | | 250–400% FPL | |
|--|---|---------|--|--------|--|-------|--|-------|
| | Change in government spending (\$ million/year) | | Change in consumer surplus (\$ million/year) | | Change in consumer surplus (\$/person-month) | | Change in consumer surplus (\$/person-month) | |
| | Bounds/Point estimate | | Bounds/Point estimate | | Bounds/Point estimate | | Bounds/Point estimate | |
| Nonparametric | –601.73 | –207.05 | –73.67 | –62.49 | –3.10 | –2.59 | –1.50 | –1.32 |
| $\kappa_{age} = 0.4$ $\kappa_{inc} = 0$ | –622.58 | –217.51 | –74.00 | –62.17 | –3.11 | –2.58 | –1.51 | –1.31 |
| $\kappa_{age} = \infty$ $\kappa_{inc} = 0$ | –750.84 | –188.13 | –75.78 | –56.74 | –3.17 | –2.39 | –1.57 | –1.15 |
| $\kappa_{age} = 0$ $\kappa_{inc} = 0.4$ | –1136.26 | –393.73 | –72.10 | –50.82 | –3.02 | –2.04 | –1.49 | –1.17 |
| $\kappa_{age} = 0$ $\kappa_{inc} = \infty$ | –2092.82 | –281.28 | –74.55 | –11.83 | –3.08 | –0.34 | –1.60 | –0.44 |
| Logit | | –295.13 | | –81.08 | | –3.44 | | –1.78 |
| Probit | | –268.52 | | –82.38 | | –3.39 | | –1.78 |
| Mixed Logit I | | –286.42 | | –83.64 | | –3.51 | | –1.88 |
| Mixed Logit II | | –339.43 | | –57.27 | | –2.32 | | –1.40 |
| Mixed Logit III | | –266.55 | | –63.27 | | –2.60 | | –1.49 |

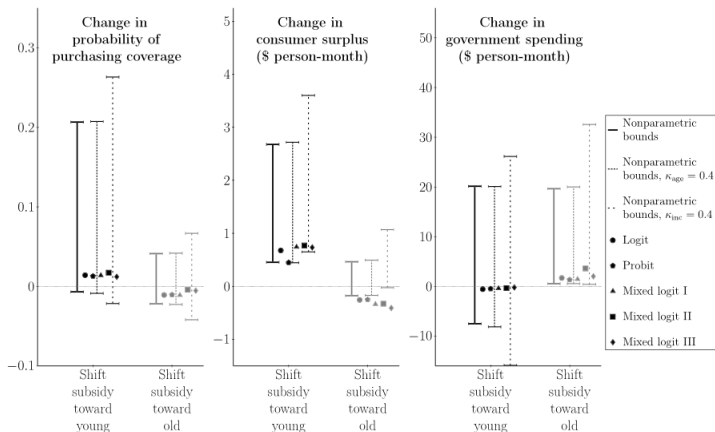
Note: Each pair of columns corresponds to a different target parameter. Lower and upper bounds are shown for the nonparametric model with different sensitivity to age and income, while single point estimates are shown for the parametric models.



(a) Counterfactual changes to premiums

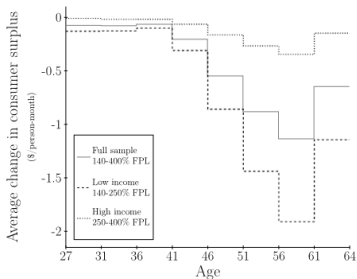


(b) Changes in enrollment by age group

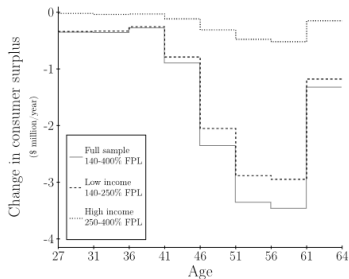


(c) Effects on aggregate outcomes

FIGURE 5.—Linking subsidies to age. *Notes:* Panel (a) illustrates the change in subsidized premiums by age under the two counterfactuals considered. Each x-axis group in panels (b) and (c) contains estimated nonparametric bounds and parametric point estimates on the indicated counterfactual, as well as the baseline value at the observed premiums.



(a) Average impacts per person/month



(b) Aggregate impacts

FIGURE 6.—Upper bounds on the change in consumer surplus from removing Silver plans. *Notes:* Each line indicates the estimated nonparametric upper bound on the change in consumer surplus for a different income group. The nonparametric lower bound is infinite.

TABLE VII
AGGREGATE IMPACTS FROM REMOVING SILVER PLANS.

| | 140–400% FPL | | 140–250% FPL | | 250–400% FPL | |
|--|--|---------|--|---------|--|--------|
| | Change in consumer surplus (\$ million/year) | | Change in consumer surplus (\$ million/year) | | Change in consumer surplus (\$ million/year) | |
| | Bounds/Point estimate | | Bounds/Point estimate | | Bounds/Point estimate | |
| Nonparametric | –∞ | –12.43 | –∞ | –10.78 | –∞ | –1.66 |
| $\kappa_{\text{age}} = 0.4 \quad \kappa_{\text{inc}} = 0$ | –∞ | –11.63 | –∞ | –10.25 | –∞ | –1.38 |
| $\kappa_{\text{age}} = \infty \quad \kappa_{\text{inc}} = 0$ | –∞ | –0.68 | –∞ | –0.68 | –∞ | –0.00 |
| $\kappa_{\text{age}} = 0 \quad \kappa_{\text{inc}} = 0.4$ | –∞ | –9.77 | –∞ | –8.36 | –∞ | –1.41 |
| $\kappa_{\text{age}} = 0 \quad \kappa_{\text{inc}} = \infty$ | –∞ | –1.71 | –∞ | –1.39 | –∞ | –0.32 |
| Logit | | –281.67 | | –248.97 | | –36.24 |
| Probit | | –290.40 | | –260.62 | | –29.79 |
| Mixed Logit I | | –292.09 | | –257.40 | | –38.29 |
| Mixed Logit II | | –148.95 | | –135.53 | | –14.48 |
| Mixed Logit III | | –173.74 | | –162.85 | | –11.65 |

Note: See notes for Table VI.

Starc (2014)

Medigap

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Saltzman
(2019)

Tebaldi,
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(2023)

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(2013)**

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Einav et al. (2013)

Starc (2014)

Medigap

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Results

Saltzman
(2019)

Tebaldi,
Torgovitsky,
and Yang
(2023)

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**Einav et al.
(2013)**

References

Einav et al. (2013)

[http://faculty.arts.ubc.ca/pschrimpf/565/
efrsc-slides_2012_05_18.pdf](http://faculty.arts.ubc.ca/pschrimpf/565/efrsc-slides_2012_05_18.pdf)

Barseghyan, Levon, Francesca Molinari, Ted O'Donoghue, and Joshua C. Teitelbaum. 2013. "The Nature of Risk Preferences: Evidence from Insurance Choices." *American Economic Review* 103 (6):2499–2529. URL <http://www.aeaweb.org/articles?id=10.1257/aer.103.6.2499>.

Bundorf, M. Kate, Jonathan Levin, and Neale Mahoney. 2012. "Pricing and Welfare in Health Plan Choice." *The American Economic Review* 102 (7):3214–3248. URL <http://www.jstor.org/stable/41724632>.

Cardon, James H. and Igal Hendel. 2001. "Asymmetric Information in Health Insurance: Evidence from the National Medical Expenditure Survey." *The RAND Journal of Economics* 32 (3):408–427. URL <http://www.jstor.org/stable/2696362>.

Deb, Rahul, Yuichi Kitamura, John K. H. Quah, and Jörg Stoye. 2021. "Revealed Price Preference: Theory and Empirical Analysis."

- Einav, Liran and Amy Finkelstein. 2011. "Selection in Insurance Markets: Theory and Empirics in Pictures." *Journal of Economic Perspectives* 25 (1):115–38. URL <http://www.aeaweb.org/articles?id=10.1257/jep.25.1.115>.
- Einav, Liran, Amy Finkelstein, and Jonathan Levin. 2010. "Beyond testing: Empirical models of insurance markets." *Annual Review of Economics* 2 (1):311–336. URL http://www.annualreviews.org/doi/10.1146/annurev.economics.050708.143254#_i14.
- Einav, Liran, Amy Finkelstein, and Neale Mahoney. 2021. "Chapter 14 - The IO of selection markets." In *Handbook of Industrial Organization, Volume 5, Handbook of Industrial Organization*, vol. 5, edited by Kate Ho, Ali Hortaçsu, and Alessandro Lizzeri. Elsevier, 389–426. URL <https://www.sciencedirect.com/science/article/pii/S1573448X21000145>.

- Einav, Liran, Amy Finkelstein, Stephen P. Ryan, Paul Schrimpf, and Mark R. Cullen. 2013. "Selection on Moral Hazard in Health Insurance." *American Economic Review* 103 (1):178–219. URL <http://www.aeaweb.org/articles?id=10.1257/aer.103.1.178>.
- Gaynor, Martin, Kate Ho, and Robert J. Town. 2015. "The Industrial Organization of Health-Care Markets." *Journal of Economic Literature* 53 (2):235–84. URL <http://www.aeaweb.org/articles?id=10.1257/jel.53.2.235>.
- Handel, Ben and Kate Ho. 2021. "Chapter 16 - The industrial organization of health care markets." In *Handbook of Industrial Organization, Volume 5, Handbook of Industrial Organization*, vol. 5, edited by Kate Ho, Ali Hortaçsu, and Alessandro Lizzeri. Elsevier, 521–614. URL <https://www.sciencedirect.com/science/article/pii/S1573448X21000169>.

- Handel, Benjamin R. 2013. "Adverse Selection and Inertia in Health Insurance Markets: When Nudging Hurts." *American Economic Review* 103 (7):2643–82. URL <http://www.aeaweb.org/articles?id=10.1257/aer.103.7.2643>.
- Handel, Benjamin R. and Jonathan T. Kolstad. 2015. "Health Insurance for "Humans": Information Frictions, Plan Choice, and Consumer Welfare." *American Economic Review* 105 (8):2449–2500. URL <http://www.aeaweb.org/articles?id=10.1257/aer.20131126>.
- Saltzman, Evan. 2019. "Demand for health insurance: Evidence from the California and Washington ACA exchanges." *Journal of Health Economics* 63:197–222. URL <https://www.sciencedirect.com/science/article/pii/S016762961730070X>.

Starc, Amanda. 2014. "Insurer pricing and consumer welfare: evidence from Medigap." *The RAND Journal of Economics* 45 (1):198–220. URL <http://dx.doi.org/10.1111/1756-2171.12048>.

Tebaldi, Pietro, Alexander Torgovitsky, and Hanbin Yang. 2023. "Nonparametric Estimates of Demand in the California Health Insurance Exchange." *Econometrica* 91 (1):107–146. URL <https://onlinelibrary.wiley.com/doi/abs/10.3982/ECTA17215>.